## STAT253/317 Winter 2021 Lecture 24

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- Brownian Motion with Drift
- Stopping Time, Strong Markov Property (Review)
- Wald's Identities for Brownian Motion

A stochastic process  $\{B(t), t \ge 0\}$  is said to be a *Brownian motion* process with drift coefficient  $\mu$  and variance parameter  $\sigma^2$  if

(i) 
$$B(0) = 0$$
;  
(ii)  $\{B(t), t \ge 0\}$  has stationary and independent increments;  
(iii) for every  $t \ge 0, s \ge 0$ ,  $B(t + s) - B(s) \sim N(\mu t, \sigma^2 t)$ 

# Stopping Time (Review)

For a continuous time stochastic process  $\{X(t), t \ge 0\}$ , a *stopping* time T with respect to  $\{X(t), t \ge 0\}$  is a nonnegative random variable, such that the event  $\{T \le t\}$  depends only on  $\{X(s), 0 \le s \le t\}$  but not  $\{X(s), s > t\}$ .

**Remark:** If T is a stopping time with respect to  $\{X(t), t \ge 0\}$ , for each non-random n > 0, the stopping time truncated at n

 $(T \land n)$  defined as min(T, n)

is also a stopping time with respective to  $\{X(t), t \ge 0\}$ . Reason:  $\{(T \land n) \le t\} = \{T \le t\} \cup \{n \le t\}$ 

- The event  $\{n \le t\}$  is non-random, does not depend on  $\{X(s)\}$
- The event {T ≤ t} depends only on {X(s), 0 ≤ s ≤ t} but not {X(s), s > t} since T is a stopping time

Hence the event  $\{(T \land n) \le t\}$  depends on  $\{X(s), 0 \le s \le t\}$  only but not  $\{X(s), s > t\}$ , which shows  $(T \land n)$  is also a stopping time. Lecture 24 - 3

#### Strong Markov Property (Review)

Let  $\{B(t), t \ge 0\}$  be a Brownian Motion (with drift  $\mu$ ), and let T be a stopping time respective to  $\{B(t), t \ge 0\}$ . Then

(a) Define 
$$Z(t) = B(t + T) - B(T)$$
,  $t \ge 0$ .

Then  $\{Z(t), t \ge 0\}$  is also a Brownian Motion with drift  $\mu$ (b) For each t > 0,  $\{Z(s), 0 \le s \le t\}$  is independent of  $\{B(s), 0 \le s \le T\}$ 

**Remark:** If T is not a stopping time, the Strong Markov Property may not be true. For example, let

$$T = T_{\max} = \min \Big\{ t : B(t) = \max_{0 \le s \le 1} B(s) \Big\},$$

where  $\{B(t), t \ge 0\}$  is a standard Brownian motion.

T<sub>max</sub> is not a stopping time since the event {T<sub>max</sub> ≤ t} depends not just {B(s), 0 ≤ s ≤ t}, but on the entire {B(s), 0 ≤ s ≤ 1}.

▶ Since  $B(T_{max})$  will be the maximum of  $\{B(s), 0 \le s \le 1\}$ ,  $B(t + T_{max}) - B(T_{max})$  will be  $\le 0$  for  $t \le 1 - T_{max}$ , and hence is not Brownian motion Lecture 24 - 4

## Wald's Identities for Brownian Motion

If  $\{B(t), t \ge 0\}$  is a Brownian motion process with drift  $\mu$  and variance parameter  $\sigma^2$ , and T is a **bounded stopping time** with respect to  $\{B(t)\}$ , then

(i) 
$$\mathbb{E}[B(T)] = \mu \mathbb{E}[T]$$
,  
(ii)  $\mathbb{E}[B^2(T)] = \sigma^2 \mathbb{E}[T] + \mu^2 \mathbb{E}[T^2]$ ,  
(iii)  $\mathbb{E}[e^{\theta B(T) - (\theta \mu + \frac{\theta^2 \sigma^2}{2})T}] = 1$  for all  $\theta \in \mathbb{R}$   
Remark:

For *nonrandom* times T = t, the identities follows from the elementary properties of the normal distribution

#### ▶ If *T* is *unbounded*, the identities may not be true

- Example: if T = T<sub>1</sub> be the hitting time to value 1 of a standard Brownian motion, then B(T) = 1. So E[B(T)] ≠ 0.
- ▶ If *T* is not a stopping time, the identities may also fail.
  - ► Example: if  $T = T_{\max} = \min\{t : B(t) = \max_{0 \le s \le 1} B(s)\}$ then  $\mathbb{E}[B(T_{\max})] = \mathbb{E}[\max_{0 \le s \le 1} B(s)] > 0.$

#### Application of Wald's Identities

For constants a, b > 0 Let  $T = T_{-a,b}$  be the first time t such that the standard Brownian Motion process hit -a or b

$$T_{-a,b} = \min\{t : B(t) = -a, \text{ or } B(t) = b\}$$

T is a stopping time since the event

$$\{T \leq t\} = \Big\{\max_{0 \leq s \leq t} B(s) \geq b\Big\} \bigcup \Big\{\min_{0 \leq s \leq t} B(s) \leq -a\Big\},$$

depends on  $\{B(s), 0 \le s \le t\}$  only.

- T is finite, but <u>unbounded</u>  $\Rightarrow$  Wald's identities may <u>not</u> apply.
- ► However, for each integer n ≥ 1, the random variable T ∧ n = min(T, n) is a bounded stopping time. By the first and second Wald's identities, we have

$$\mathbb{E}[B(T \wedge n)] = 0$$
 and  $\mathbb{E}[B^2(T \wedge n)] = \mathbb{E}[T \wedge n]$ 

Application of Wald's Identities (Cont'd)

- From that −a ≤ B(T ∧ n) ≤ b, we know |B(T ∧ n)| is uniformly bounded by a + b for all n
- As  $P(T < \infty) = 1$ , we have  $\lim_{n \to \infty} B(T \land n) = B(T)$  w/ prob. 1.
- By Bounded Convergence Theorem,

$$\mathbb{E}[B(T)] = \lim_{n \to \infty} \mathbb{E}[B(T \land n)] = 0$$
(1)

$$\mathbb{E}[B^{2}(T)] = \lim_{n \to \infty} \mathbb{E}[B^{2}(T \wedge n)] = \lim_{n \to \infty} \mathbb{E}[T \wedge n] = \mathbb{E}[T]$$
(2)

• Because 
$$B(T) = -a$$
 or b, from that

$$\mathbb{E}[B(T)] = -a\mathrm{P}(B(T) = -a) + b\mathrm{P}(B(T) = b) = 0$$

and that  $\mathrm{P}(B(\mathcal{T})=-a)+\mathrm{P}(B(\mathcal{T})=b)=1$ , it follows that

$$P(B(T) = -a) = \frac{b}{a+b}, \quad P(B(T) = b) = \frac{a}{a+b}$$

From the above and (2), one may easily deduce that E[T] = E[B<sup>2</sup>(T)] = a<sup>2</sup>P(B(T) = −a)+b<sup>2</sup>P(B(T) = b) = ab Lecture 24 - 7

### Exercise 10.22: $T_{-a,b}$ for Brownian with Drift

Let  $\{B(t), t \ge 0\}$  be Brownian Motion with drift coefficient  $\mu \ne 0$ and variance parameter  $\sigma^2$ . For constants a, b > 0 let

$$T = T_{-a,b} = \min\{t : B(t) = -a, \text{ or } B(t) = b\}$$

T is again a finite but <u>unbounded</u> stopping time, so Wald's identities may <u>not</u> be applied directly. However, using the truncated stopping time  $T \wedge n = \min(T, n)$  and Bounded Convergence Theorem, we can prove that the first Wald's identity holds for T

$$\mu \mathbb{E}[T] = \mathbb{E}[B(T)] = -a P(B(T) = -a) + b P(B(T) = b).$$

However, when  $\mu \neq 0$ , we cannot use this equation and that P(B(T) = -a) + P(B(T) = b) = 1 to solve for P(B(T) = -a) and P(B(T) = b) since  $\mathbb{E}[T]$  is unknown. Instead we will use the third Wald's identity.

# Exercise 10.22: $T_{-a,b}$ for Brownian with Drift (Cont'd)

By the third Wald's identity, we have

$$\mathbb{E}[e^{\theta B(T \wedge n) - (\theta \mu + \frac{\theta^2 \sigma^2}{2})(T \wedge n)}] = 1 \quad \text{for all } \theta \in \mathbb{R}.$$
(3)

Let us choose  $\theta = \theta_0 = -2\mu/\sigma^2$  so that the 2nd term in the exponent of (3) vanishes. So

$$\mathbb{E}[e^{\theta_0 B(T \wedge n)}] = 1$$

$$\ -a \le B(T \land n) \le b \Rightarrow |B(T \land n)| \le a + b \\ \Rightarrow e^{\theta_0 B(T \land n)} \le e^{\theta_0 (a+b)}$$

By the Bounded Convergence Theorem,

$$1 = \lim_{n \to \infty} \mathbb{E}[e^{\theta_0 B(T \land n)}] = \mathbb{E}[e^{\theta_0 B(T)}]$$
$$= e^{-\theta_0 a} \mathcal{P}(B(T) = -a) + e^{\theta_0 b} \mathcal{P}(B(T) = b)$$

#### Exercise 10.22: $T_{-a,b}$ for Brownian with Drift (Cont'd)

Solving the equation

$$1 = e^{-\theta_0 a} \mathbf{P}(B(T) = -a) + e^{\theta_0 b} \mathbf{P}(B(T) = b)$$

and the equation P(B(T) = -a) + P(B(T) = b) = 1 for P(B(T) = -a) and P(B(T) = b), one can get that

$$P(B(T) = -a) = \frac{1 - e^{\theta_0 b}}{e^{-\theta_0 a} - e^{\theta_0 b}}, \quad P(B(T) = b) = \frac{e^{-\theta_0 a} - 1}{e^{-\theta_0 a} - e^{\theta_0 b}}$$

**Theorem 1**. Let  $\{B(t), t \ge 0\}$  be a Brownian Motion with drift coefficient  $\mu \ne 0$  and variance parameter  $\sigma^2$ , the probability that the process reach b > 0 before hitting -a < 0 is given by

$$P(B(T_{-a,b}) = b) = \frac{\exp(2\mu a/\sigma^2) - 1}{\exp(2\mu a/\sigma^2) - \exp(-2\mu b/\sigma^2)}$$

### Proof of Wald's Identities for Brownian Motion

- Since T is bounded, there is a nonrandom  $N < \infty$  such that P(T < N) = 1
- By the Strong Markov Property, the post-T process B(t + T) B(T) is
  - ▶ also a Brownian Motion process with drift  $\mu$  and variance parameter  $\sigma^2$ , and
  - ▶ independent of {B(s), 0 ≤ s ≤ T}, and in particular, independent of the random vector (T, B(T)).
- Hence, given that T = s the conditional distribution of B(N) − B(T) is normal with mean µ(N − s) and variance σ<sup>2</sup>(N − s). It follows that

$$\mathbb{E}\left[e^{\theta[B(N)-B(T)]-\theta\mu(N-T)-\frac{\theta^2\sigma^2(N-T)}{2}}\Big|T,B(T)\right]=1$$

# Proof of Wald's Identities (Cont'd)

Therefore

$$\begin{split} \mathbb{E}[e^{\theta B(T) - \theta \mu T - \frac{\theta^2 \sigma^2 T}{2}}] &= \mathbb{E}[e^{\theta B(T) - \theta \mu T - \frac{\theta^2 \sigma^2 T}{2}}] \times 1 \\ &= \mathbb{E}[e^{\theta B(T) - \theta \mu T - \frac{\theta^2 \sigma^2 T}{2}}] \\ &\times \mathbb{E}\Big[e^{\theta [B(N) - B(T)] - \theta \mu (N - T) - \frac{\theta^2 \sigma^2 (N - T)}{2}} \Big| T, B(T)\Big] \\ &= \mathbb{E}\Big[\mathbb{E}\Big[e^{\theta B(T) - \theta \mu T - \frac{\theta^2 \sigma^2 T}{2} + \theta [B(N) - B(T)] - \theta \mu (N - T) - \frac{\theta^2 \sigma^2 (N - T)}{2}} \Big| T, B(T)\Big]\Big] \\ &= \mathbb{E}\Big[\mathbb{E}\Big[e^{\theta B(N) - \theta \mu N - \frac{\theta^2 \sigma^2 N}{2}} \Big| T, B(T)\Big]\Big] \\ &= \mathbb{E}[e^{\theta B(N) - \theta \mu N - \frac{\theta^2 \sigma^2 N}{2}}] = 1 \end{split}$$

This proves the third identity.

The first and second identity can be derived by differentiating the third identity with respective to  $\theta$  once and twice respectively, and letting  $\theta = 0$ .