

Multi-way Contingency Tables

Outline

- ▶ Multiway Contingency Tables
 - ▶ Flat Contingency Tables
 - ▶ How to effectively convey information in multiway contingency tables?
 - ▶ Manipulation of 'Flat' Contingency Tables in R
- ▶ Logistic Models for Multi-way Contingency Tables

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A better table:

	<i>Muscle</i>	<i>Tension</i>			
		Type 1		Type 2	
<i>Weight</i>	<i>Drug</i>	High	Low	High	Low
High	1	3	3	23	41
	2	21	10	11	21
Low	1	22	45	4	6
	2	32	23	12	22

Conditional odds ratios between Drug and Tension:

<i>Wt.</i>	<i>Muscle</i>	
	Type 1	Type 2
High	$\frac{3 \times 10}{3 \times 21} \approx 0.48$	$\frac{23 \times 21}{41 \times 11} \approx 1.07$
Low	$\frac{22 \times 23}{45 \times 32} \approx 0.35$	$\frac{4 \times 22}{6 \times 12} \approx 1.22$

- ▶ The table splits in to four partial tables for the primary predictor (Drug) and the response (Tension), controlling for the other two variables.
- ▶ Conditional odds ratios between Drug and Tension can be easily computed from this table but not from the table on the previous slide.
- ▶ Tip: response and the primary predictor (if any) should be placed in the inner most layer of the table

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Example (Mouse Muscle Tension)

A study to examine relationship between two drugs and muscle tension

Response: Tension — change in muscle tension: High, Low

Explanatory variables

- ▶ Drug: drug 1, drug (primary)
- ▶ Weight: weight of muscle: High, Low
- ▶ Muscle: muscle type: 1, 2

A four-way "flat" contingency table ($2 \times 2 \times 2 \times 2$):

<i>Tension</i>	<i>Weight</i>	<i>Drug 1</i>		<i>Drug 2</i>	
		1	2	1	2
High	High	3	23	21	11
	Low	22	4	32	12
Low	High	3	41	10	21
	Low	45	6	23	22

This flat table is bad because ...

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Example (Mouse Muscle Tension)

Conditional distributions of Tension given Drug, Weight, and Muscle type:

<i>Weight</i>	<i>Drug</i>	<i>Muscle</i>			
		Type 1		Type 2	
		High	Low	High	Low
High	1	50%	50%	36%	64%
	2	68%	32%	34%	66%
Low	1	33%	67%	40%	60%
	2	58%	42%	35%	65%

Observation:

- ▶ For Type 1 muscle, Drug 1 looks more effective in lowering muscle tension than Drug 2 does
- ▶ For Type 2 muscle, the effect of the two drugs looks similar

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Another Example of “Flat” Contingency Tables

Example (Titanic)

Four-way table: $(2 \times 2 \times 4 \times 2)$

Breakup of people on Titanic by Sex, Age, Class, and Survival

Class	Sex	Adult				Child			
		Age	Female		Male		Female		Male
	Survived		No	Yes	No	Yes	No	Yes	No
1st		4	140	118	57	0	1	0	5
2nd		13	80	154	14	0	13	0	11
3rd		89	76	387	75	17	14	35	13
Crew		3	20	670	192	0	0	0	0

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```
> data.frame(weight, muscle, drug, tension, Freq)
  weight muscle drug tension Freq
1   High     1    1   High     3
2   High     1    2   High    21
3   High     2    1   High    23
4   High     2    2   High    11
5   Low      1    1   High    22
6   Low      1    2   High    32
7   Low      2    1   High     4
8   Low      2    2   High    12
9   High     1    1   Low     3
10  High     1    2   Low    10
11  High     2    1   Low    41
12  High     2    2   Low    21
13  Low      1    1   Low    45
14  Low      1    2   Low    23
15  Low      2    1   Low     6
16  Low      2    2   Low    22
```

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Manipulating Flat Contingency Tables in R

```
> mouse.muscle = read.table("mousemuscle.dat",header=T)
> mouse.muscle
      W M D tension.high tension.low
1 High 1 1           3           3
2 High 1 2           21          10
3 High 2 1           23          41
4 High 2 2           11          21
5 Low  1 1           22          45
6 Low  1 2           32          23
7 Low  2 1            4           6
8 Low  2 2           12          22

> attach(mouse.muscle)
> Freq = c(tension.high,tension.low)
> weight = rep(W,2)
> muscle = rep(M,2)
> drug = rep(D,2)
> tension = c(rep("High",8),rep("Low",8))
```

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`xtabs` creates multiway tables in R, but the output is awkward.

```
> muscle.tab = xtabs(Freq ~ weight + muscle + drug + tension)
> muscle.tab
, , drug = 1, tension = High
      muscle
weight 1 2
      High 3 23
      Low  22  4

, , drug = 2, tension = High
      muscle
weight 1 2
      High 21 11
      Low  32  12

, , drug = 1, tension = Low
      muscle
weight 1 2
      High 3 41
      Low  45  6

, , drug = 2, tension = Low
      muscle
weight 1 2
      High 10 21
      Low  23  22
```

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`fable()` can print flat multi-way tables. Row and column variables in a flat table can be specified using `row.vars` and `col.vars`, and they are ordered from outer to inner layers.

```
> ftable(muscle.tab, row.vars=c("weight","drug"),
         col.vars=c("muscle","tension"))
      muscle  1      2
      tension High Low High Low
weight drug
High  1      3   3  23  41
      2      21  10  11  21
Low   1      22  45   4   6
      2      32  23  12  22

> ftable(muscle.tab, row.vars=c("tension","weight"),
         col.vars=c("drug","muscle"))
      drug  1      2
      muscle 1  2  1  2
tension weight
High   High      3 23 21 11
      Low      22  4 32 12
Low   High      3 41 10 21
      Low      45  6 23 22
```

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Flat Marginal Tables

If a variable in the contingency table is neither specified in `row.vars` nor in `col.vars`, then it is ignored. The output is a marginal table of the specified variables, e.g., the following is a marginal table for `drug`, `muscle`, and `tension`, ignoring `weight`:

```
> ftable(muscle.tab,row.vars="drug", col.vars=c("muscle","tension"))
      muscle  1      2
      tension High Low High Low
drug
1      25  48  27  47
2      53  33  23  43
```

Marginal table for `drug` and `tension`, ignoring `weight` and `muscle`:

```
> ftable(muscle.tab,row.vars="drug", col.vars="tension")
      tension High Low
drug
1      52  95
2      76  76
```

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Flat Table for Conditional Distributions

First compute the conditional distributions of `tension` given the rest using `prop.table`, and then print it using `fable`.

```
> muscle.p.tab = prop.table(muscle.tab,1:3)
> ftable(muscle.p.tab,row.vars=c("weight","drug"),
         col.vars=c("muscle","tension"))
      muscle  1      2
      tension High Low High Low
weight drug
High  1      0.5000000 0.5000000 0.3593750 0.6406250
      2      0.6774194 0.3225806 0.3437500 0.6562500
Low   1      0.3283582 0.6716418 0.4000000 0.6000000
      2      0.5818182 0.4181818 0.3529412 0.6470588

> round(100*ftable(muscle.p.tab,row.vars=c("weight","drug")))
      muscle  1      2
      tension High Low High Low
weight drug
High  1      50  50  36  64
      2      68  32  34  66
Low   1      33  67  40  60
      2      58  42  35  65
```

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Conditional Distributions in Marginal Tables

Say, we want the condition distribution of `tension` given `drug` and `muscle`, ignoring `weight`.

1. Create the marginal table, with "response" as the only column variable, and other var. in the marginal table as the row var..

```
> temp1 = ftable(muscle.tab, row.vars=c("muscle","drug"),
                col.vars="tension")
> temp1
      tension High Low
muscle drug
1      1      25  48
      2      53  33
2      1      27  47
      2      23  43
```

2. Find the conditional distribution using `prop.table`.

```
> temp2 = prop.table(temp1, 1)
> temp2
      tension High Low
muscle drug
1      1      0.3424658 0.6575342
      2      0.6162791 0.3837209
2      1      0.3648649 0.6351351
      2      0.3484848 0.6515152
```

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Conditional Distributions in Marginal Tables

3 If necessary, reshape the condition distribution table using `fTable`

```
> fTable(temp2, row.vars="drug", col.vars=c("muscle", "tension"))
      muscle      1      2
      tension  High  Low  High  Low
drug
1          0.3424658 0.6575342 0.3648649 0.6351351
2          0.6162791 0.3837209 0.3484848 0.6515152

# converted into percentages
> round(100*fTable(temp2, row.vars="drug",
                  col.vars=c("muscle", "tension")))
      muscle      1      2
      tension  High  Low  High  Low
drug
1          34   66   36   64
2          62   38   35   65
```

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Let

- ▶ A_i , $i = 1, \dots, a$ be the dummy variables for levels of A
- ▶ B_j , $j = 1, \dots, b$ be the dummy variables for levels of B
- ▶ C_k , $k = 1, \dots, c$ be the dummy variables for levels of C

The model formula

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C + \beta_{ij}^{AB} + \beta_{jk}^{BC} + \beta_{ik}^{AC} + \beta_{ijk}^{ABC}$$

can be written in terms of the dummy variables as

$$\begin{aligned} \text{logit}(\pi_{ijk}) = & \alpha + \sum_{\ell=1}^a \beta_{\ell}^A A_{\ell} + \sum_{m=1}^b \beta_m^B B_m + \sum_{n=1}^c \beta_n^C C_n \\ & + \sum_{\ell=1}^a \sum_{m=1}^b \beta_{\ell m}^{AB} A_{\ell} B_m + \sum_{m=1}^b \sum_{n=1}^c \beta_{mn}^{BC} B_m C_n + \sum_{\ell=1}^a \sum_{n=1}^c \beta_{\ell n}^{AC} A_{\ell} C_n \\ & + \sum_{\ell=1}^a \sum_{m=1}^b \sum_{n=1}^c \beta_{\ell mn}^{ABC} A_{\ell} B_m C_n \end{aligned}$$

- ▶ How many parameters are there?

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Logistic Models for Multi-way Contingency Tables

Let's start w/ models for 4-way tables (1 response + 3 predictors)

- ▶ categorical predictors: A, B, C , with a, b, c levels respectively
- ▶ response: $Y = 0$ or 1

Let

$$\pi_{ijk} = P(Y = 1 | A = i, B = j, C = k)$$

The most complex model for a 4-way table is the **three way interaction model**, denoted as $A * B * C$, including all main effects and 2-way, 3-way interactions

$$A + B + C + A * B + B * C + A * C + A * B * C$$

The model formula is

$$\text{logit}(\pi_{ijk}) = \alpha + \underbrace{\beta_i^A + \beta_j^B + \beta_k^C}_{\text{main effects}} + \underbrace{\beta_{ij}^{AB} + \beta_{jk}^{BC} + \beta_{ik}^{AC}}_{\text{two-way interactions}} + \beta_{ijk}^{ABC}$$

for $i = 1, \dots, a, j = 1, \dots, b, k = 1, \dots, c$.

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In the 3-way on the previous page, many parameters are redundant because

$$A_1 + \dots + A_a = 1, \quad B_1 + \dots + B_b = 1, \quad C_1 + \dots + C_c = 1.$$

So, need constraints on the parameters.

Some commonly used constraints are

- ▶ Baseline constraints:

$$\begin{aligned} \beta_1^A = \beta_1^B = \beta_1^C &= 0 \\ \beta_{1j}^{AB} = \beta_{i1}^{AB} = \beta_{1k}^{BC} = \beta_{j1}^{BC} = \beta_{1k}^{AC} = \beta_{i1}^{AC} &= 0 \\ \beta_{1jk}^{ABC} = \beta_{i1k}^{ABC} = \beta_{ij1}^{ABC} &= 0 \end{aligned}$$

- ▶ Sum-to-zero constraints:

$$\begin{aligned} \sum_{i=1}^a \beta_1^A = \sum_{j=1}^b \beta_1^B = \sum_{k=1}^c \beta_1^C &= 0 \\ \sum_{i=1}^a \beta_{ij}^{AB} = \sum_{j=1}^b \beta_{ij}^{AB} = \sum_{j=1}^b \beta_{jk}^{BC} = \sum_{k=1}^c \beta_{jk}^{BC} = \sum_{i=1}^a \beta_{ik}^{AC} = \sum_{k=1}^c \beta_{ik}^{AC} &= 0 \\ \sum_{i=1}^a \beta_{1jk}^{ABC} = \sum_{j=1}^b \beta_{ijk}^{ABC} = \sum_{k=1}^c \beta_{ijk}^{ABC} &= 0 \end{aligned}$$

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Regardless of the constraints used,

- ▶ the effective # of parameters for a main effect is
number of levels – 1
- ▶ the effective # of parameters for an interaction is the product of (number of levels – 1) for each factor involved in the interaction.

The total number of effective parameters is

$$\begin{aligned}
 & 1 + \underbrace{(a-1)}_{\text{A main effects}} + \underbrace{(b-1)}_{\text{B main effects}} + \underbrace{(c-1)}_{\text{C main effects}} \\
 & + \underbrace{(a-1)(b-1)}_{\text{AB interactions}} + \underbrace{(b-1)(c-1)}_{\text{BC interactions}} + \underbrace{(a-1)(c-1)}_{\text{AC interactions}} \\
 & + \underbrace{(a-1)(b-1)(c-1)}_{\text{ABC interactions}} \\
 & = abc
 \end{aligned}$$

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- ▶ In all the models, constraints on the main effects and interactions are the same as those on the corresponding parameters in the 3-way interaction models.
- ▶ Generally, models must maintain *hierarchy* — cannot include an interaction terms without including the relevant main effects and lower order interactions

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There are several simplifications of the 3-way interaction model, such as

- ▶ Model $A * B + B * C + A * C$:

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C + \beta_{ij}^{AB} + \beta_{jk}^{BC} + \beta_{ik}^{AC}$$

- ▶ Model $A * B + A * C$

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C + \beta_{ij}^{AB} + \beta_{ik}^{AC}$$

- ▶ Model $A + B * C$:

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C + \beta_{jk}^{BC}$$

- ▶ Model $A + B + C$:

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C$$

- ▶ Model $A * B$:

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_{ij}^{AB}$$

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Interpretation of Model $A + B + C$ and its Coefficients

In Model $A + B + C$:

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C$$

- ▶ There are $1 + (a-1) + (b-1) + (c-1)$ effective parameters.
- ▶ Under the baseline constraint $\beta_1^A = \beta_1^B = \beta_1^C = 0$,
 - ▶ The odds of $Y = 1$ when $A = i$ is the odds of $Y = 1$ when $A = 1$ multiplied by a factor of $e^{\beta_i^A}$, regardless of B and C
 - ▶ Interpretation for $e^{\beta_j^B}$ and $e^{\beta_k^C}$: Ditto
- ▶ Homogeneous YA , YB , and YC association,

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Homogeneous Association Revisit

In a 3-way table, if XY has homogeneous association given Z , then so do YZ given X and XZ given Y .

	$Z = 1$		$Z = 2$	
	$X = 1$	$X = 2$	$X = 1$	$X = 2$
$Y = 1$	a	b	A	B
$Y = 2$	c	d	C	D

Homogeneous XY association given Z means

$$\theta_{XY(1)} = \frac{ad}{cb} = \frac{AD}{CB} = \theta_{XY(2)}$$

$$\iff \theta_{YZ(1)} = \frac{aC}{cA} = \frac{bD}{dB} = \theta_{YZ(2)}$$

which means homogeneous YZ association given X .

	$X = 1$		$X = 2$	
	$Z = 1$	$Z = 2$	$Z = 1$	$Z = 2$
$Y = 1$	a	A	b	B
$Y = 2$	c	C	d	D

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Interpretation of Model $A * B$ and its Coefficients

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_{ij}^{AB}$$

- ▶ Number of parameters: $1 + (a - 1) + (b - 1) + (a - 1)(b - 1)$
- ▶ C is not in the model
 $\Rightarrow Y$ and C are conditionally independent given A, B
- ▶ Under the baseline constraint $\beta_1^A = \beta_1^B = \beta_{1j}^{AB} = \beta_{i1}^{AB} = 0$,

the conditional odds ratio for

	$Y = 1$	$Y = 0$
$A = i$		
$A = 1$		

equals

$$\begin{cases} e^{\beta_i^A} & \text{when } B = 1; \\ e^{\beta_i^A + \beta_{ij}^{AB}} & \text{when } B = j \end{cases}$$

So $e^{\beta_{ij}^{AB}}$ is a **ratio of two odds ratios**.

- ▶ Odds ratios of YA change with B (but not C),
 \Rightarrow no homogeneous YA association
- ▶ Likewise, can show that odds ratios of YB changes with A (but not C). \Rightarrow No homogeneous YB association

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Interpretation of Model $A * B * C$ and its Coefficients

$$\text{logit}(\pi_{ijk}) = \alpha + \beta_i^A + \beta_j^B + \beta_k^C + \beta_{ij}^{AB} + \beta_{jk}^{BC} + \beta_{ik}^{AC} + \beta_{ijk}^{ABC}$$

Under the baseline constraint, the conditional odds ratio for

	$Y = 1$	$Y = 0$
$A = i$		
$A = 1$		

$$\text{equals } \begin{cases} e^{\beta_i^A} & \text{when } B = 1, C = 1; \\ e^{\beta_i^A + \beta_{ij}^{AB}} & \text{when } B = j, C = 1; \\ e^{\beta_i^A + \beta_{ik}^{AC}} & \text{when } B = 1, C = k; \\ e^{\beta_i^A + \beta_{ij}^{AB} + \beta_{ik}^{AC} + \beta_{ijk}^{ABC}} & \text{when } B = j, C = k. \end{cases}$$

$$\text{So } \frac{\text{YA odds ratio when } B = j}{\text{YA odds ratio when } B = 1} = \begin{cases} e^{\beta_{ij}^{AB}} & \text{when } C = 1; \\ e^{\beta_{ij}^{AB} + \beta_{ijk}^{ABC}} & \text{when } C = k. \end{cases}$$

The 3-way interaction $e^{\beta_{ijk}^{ABC}}$ is **the ratio of the ratios of odds ratios**.

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Model $A * B + B * C + A * C$:

- ▶ YA odds ratios change with both B and C
- ▶ YB odds ratios change with both A and C
- ▶ YC odds ratios change with both A and B
- ▶ no 3-way interactions means that

$$\frac{\text{YA odds ratio when } B = j_1}{\text{YA odds ratio when } B = j_2}$$

do not change with C

Model $A * B + B * C$:

- ▶ YA odds ratios change with B but not C
- ▶ YC odds ratios change with A but not B
- ▶ YB odds ratios change with both A and C

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