

STAT 24400 Lecture 18

P-values

Tests & Confidence Intervals for Normal Distributions

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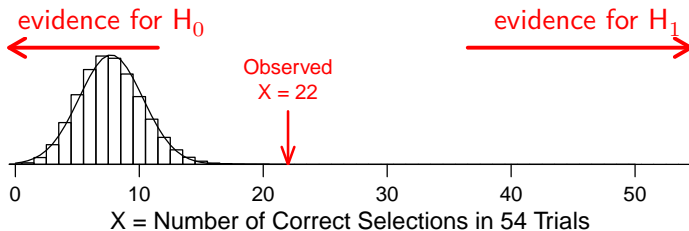
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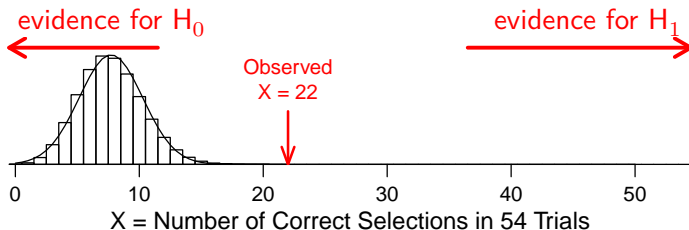
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- ▶ The P -value is a **probability**, and thus it's between 0 and 1
- ▶ This probability is calculated **assuming the H_0 is true**.
- ▶ To determine the P -value, we must first decide which values of the test statistic are the evidence for H_1 to be stronger than or as as the value obtained from our sample

P -Value — Dogs-Smell-Cancer Study



P-Value — Dogs-Smell-Cancer Study



- ▶ Observed $X = 22$
- ▶ Evidence for H_1 is stronger than or as strong as the observed $X = 22$ if $X \geq 22$
- ▶ Under H_0 , $X \sim \text{Bin}(n = 54, p = 1/7)$

$$P\text{-value} = P(X \geq 22 \mid H_0) = \sum_{k=22}^{54} \binom{54}{k} \left(\frac{1}{7}\right)^k \left(\frac{6}{7}\right)^{54-k} \approx 1.86 \times 10^{-6}$$

- ▶ Note P -value is NOT $P(X = 22 \mid H_0)$

Test Procedure Based on the P -value

As an alternative to test procedures based on rejection regions, one can use test procedures based on P -values

1. Select a significance level $\alpha = \text{the desired P(Type I error)}$.
2. Then
 - ▶ reject H_0 if the P -value $\leq \alpha$
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Remark: Hypothesis tests using a “Rejection Region” and those using the “ P -value” are equivalent. In fact,

- ▶ the test statistic is in the rejection region with significance level α if and only if the P -value $<$ the significance level α

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In the rest of L18, we will outline the test procedures for 6 major tests about the normal distribution, using both the critical-value and the P -value approach.

Six Tests for Normal Distributions

One sample: $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

1. One sample test for mean, with known σ^2
 $H_0: \mu = \mu_0$ v.s. $H_1: \mu \neq \mu_0$ (or $\mu > \mu_0, \mu < \mu_0$)
2. One sample test for mean, with unknown σ^2
 $H_0: \mu = \mu_0$ v.s. $H_1: \mu \neq \mu_0$ (or $\mu > \mu_0, \mu < \mu_0$)
3. One sample test for variance, with unknown μ
 $H_0: \sigma^2 = \sigma_0^2$ v.s. $H_1: \sigma^2 \neq \sigma_0^2$ ($\sigma^2 > \sigma_0^2, \sigma^2 < \sigma_0^2$)

Two indep samples: $X_{11}, \dots, X_{1n_1} \stackrel{\text{iid}}{\sim} N(\mu_1, \sigma_1^2)$, and $X_{21}, \dots, X_{2n_2} \stackrel{\text{iid}}{\sim} N(\mu_2, \sigma_2^2)$

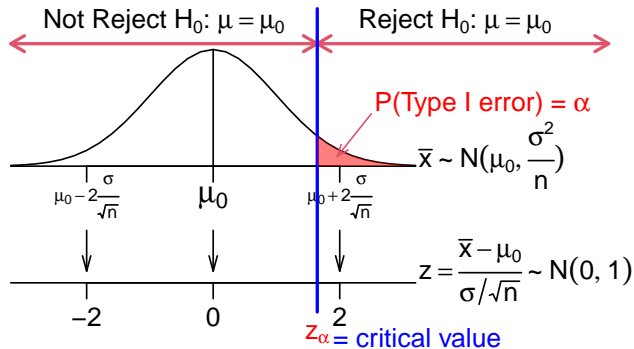
4. Two sample tests for mean, assuming $\sigma_1^2 = \sigma_2^2$
 $H_0: \mu_1 = \mu_2$ v.s. $H_1: \mu_1 \neq \mu_2$ (or $\mu_1 > \mu_2, \mu_1 < \mu_2$)
5. Two sample tests for mean, NOT assuming $\sigma_1^2 = \sigma_2^2$
 $H_0: \mu_1 = \mu_2$ v.s. $H_1: \mu_1 \neq \mu_2$ (or $\mu_1 > \mu_2, \mu_1 < \mu_2$)
6. Two sample tests for variance, μ_1 and μ_2 unknown
 $H_0: \sigma_1^2 = \sigma_2^2$ v.s. $H_1: \sigma_1^2 \neq \sigma_2^2$ (or $\sigma_1^2 > \sigma_2^2, \sigma_1^2 < \sigma_2^2$)

One Sample Tests for Mean, Known σ^2

Upper One-Sided One Sample Tests for Mean, Known σ^2

The test statistic for testing $H_0: \mu = \mu_0$ against $H_1: \mu > \mu_0$ is

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1), \quad \text{under } H_0: \mu = \mu_0.$$



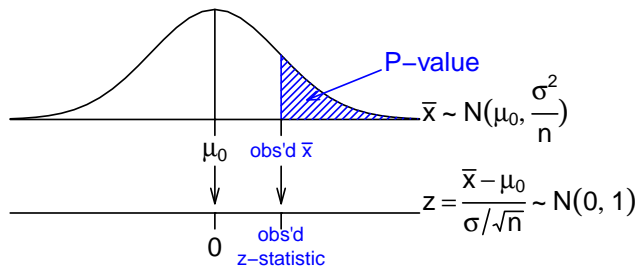
To control $P(\text{Type I error}) = P(\text{rejecting } H_0 \mid H_0 \text{ is true})$ at the significance level α , we reject H_0 when

P-value for Upper One-Sided Test

Let \bar{x} be the observed value of \bar{X} . The P -value for testing $H_0: \mu = \mu_0$ against $H_1: \mu > \mu_0$ is

$$P(Z > z) = 1 - \Phi(z), \quad \text{where } z = \text{obs'd z-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$$

or the blue shaded region below.

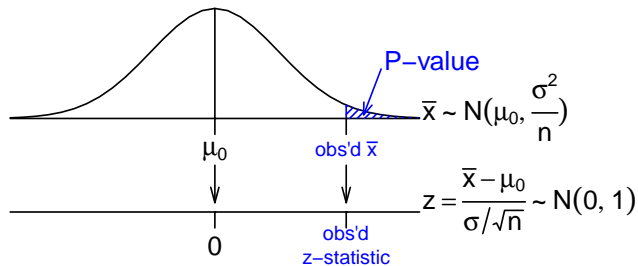


P-value for Upper One-Sided Test

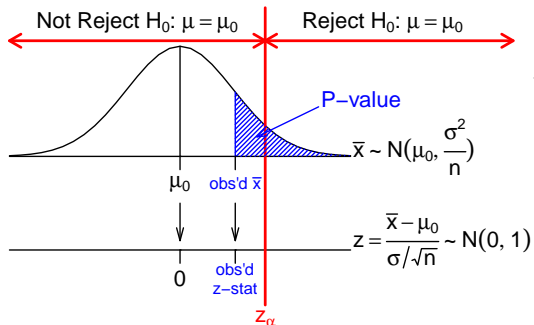
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P-value v.s. Critical Value



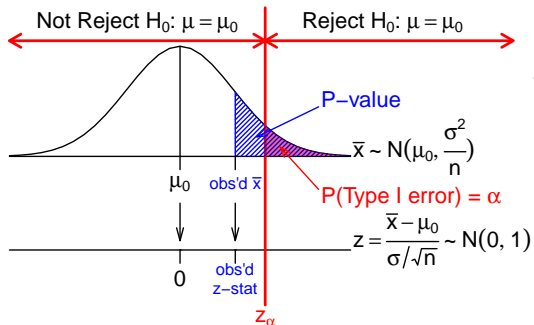
Observed that

if $z\text{-statistic} < z_\alpha$ then $P\text{-value} > \alpha$

Two equivalent approaches to test H_0 :
 $\mu = \mu_0$ v.s. $H_1: \mu > \mu_0$ and control
 $P(\text{Type I error})$ at a significance level α :

- ▶ **Critical value approach**: compute the $z\text{-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$ and the critical value z_α , and reject H_0 if the $z\text{-stat} > z_\alpha$.
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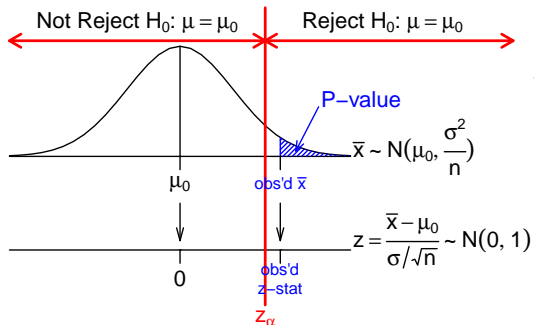
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- **Critical value approach:** compute the $z\text{-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$ and the critical value z_α , and reject H_0 if the $z\text{-stat} > z_\alpha$.
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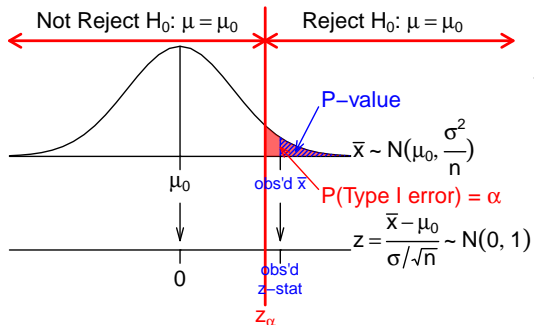
if z -statistic $< z_\alpha$ then $P\text{-value} > \alpha$

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- **P-value approach:** compute the $P\text{-value}$ from the $z\text{-stat}$ and reject H_0 when $P\text{-value} < \alpha$

P -value is the Smallest Significance Level to Reject H_0

The P -value is the **smallest significance level α at which the H_0 can be rejected**.

► e.g., the P -value for the dog study is 1.86×10^{-6} .

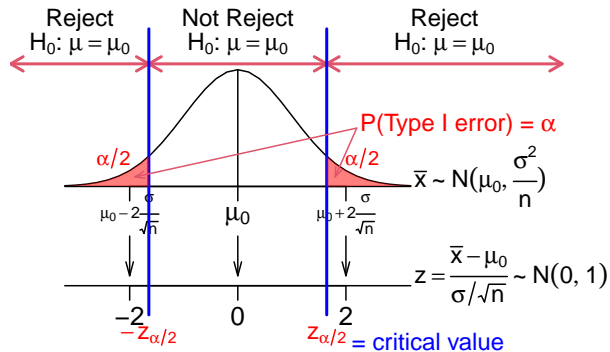
The H_0 won't be rejected unless the significance level is as small as 1.86×10^{-6}

Because of this, the P -value is alternatively referred to as the *observed significance level* for the data.

Two-Sided One Sample Tests for Mean, Known σ^2 ,

For a two-sided test of $H_0: \mu = \mu_0$ against $H_1: \mu \neq \mu_0$, the test statistic remains to be

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1), \quad \text{under } H_0: \mu = \mu_0.$$



To control $P(\text{Type I error})$ at the significance level α , reject H_0 when

$$|z\text{-stat}| = \left| \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \right| > z_{\alpha/2},$$

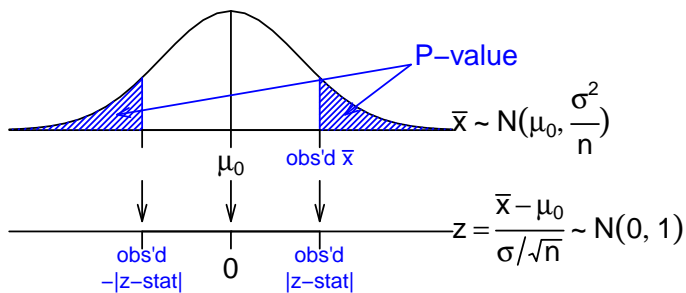
where $\Phi(z_{\alpha/2}) = 1 - \frac{\alpha}{2}$.

P-values for Two-Sided Hypothesis Tests

To test $H_0: \mu = \mu_0$ against **two-sided alternative** $H_1: \mu \neq \mu_0$, the P -value is the two-tail probability

$$P(|Z| > |z|) = 2(1 - \Phi(|z|)), \quad \text{where } z = \text{obs'd z-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$$

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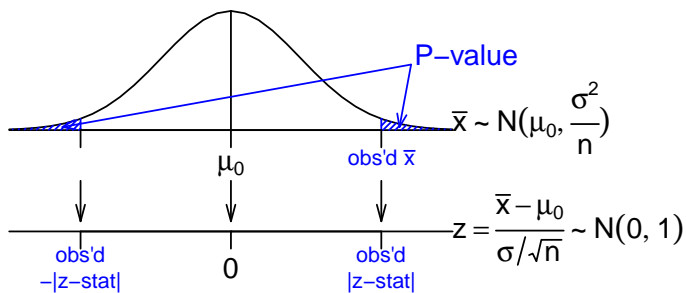


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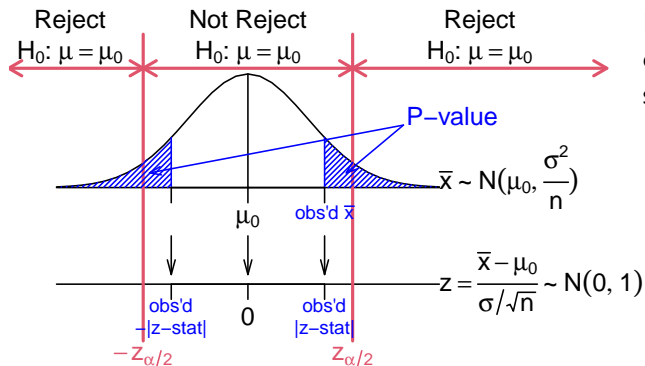
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P-value and Critical Value Approaches for Two-Sided Tests

Observed

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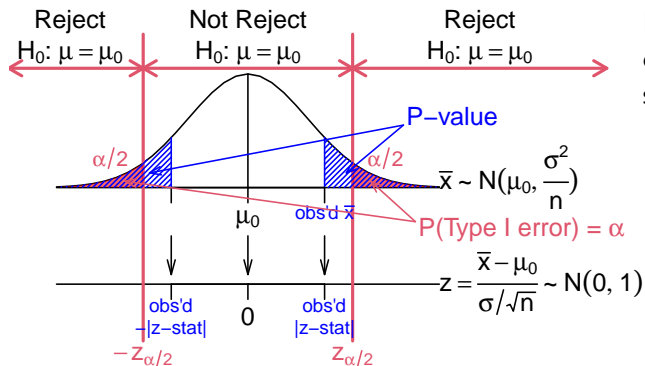
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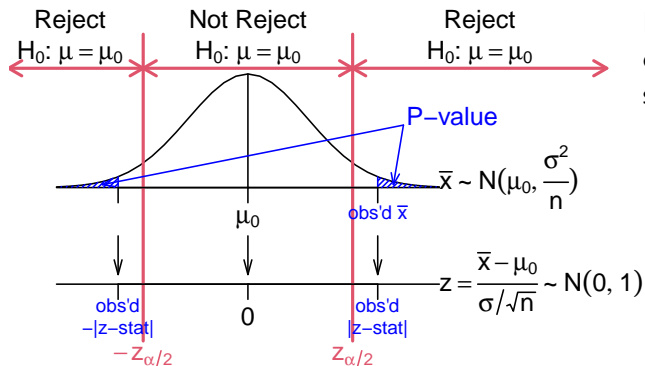
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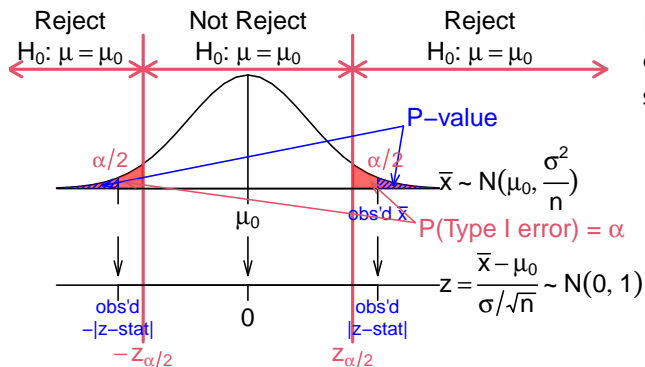
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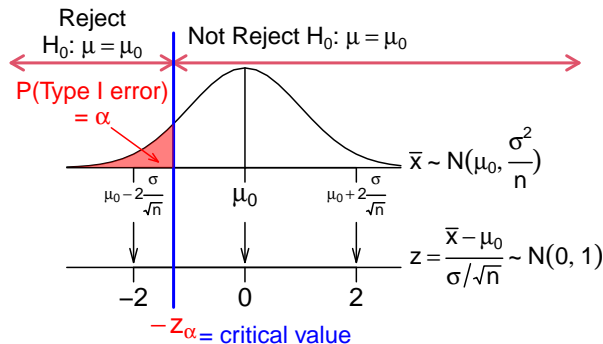
Three Types of Alternative Hypotheses:

- ▶ Upper one-sided: $H_1: \mu > \mu_0$
- ▶ Lower one-sided: $H_1: \mu < \mu_0$
- ▶ Two-sided: $H_1: \mu \neq \mu_0$

Lower One-Sided Tests

To test $H_0: \mu = \mu_0$ against the **lower one-sided** alternative $H_1: \mu < \mu_0$, the test statistic remains to be

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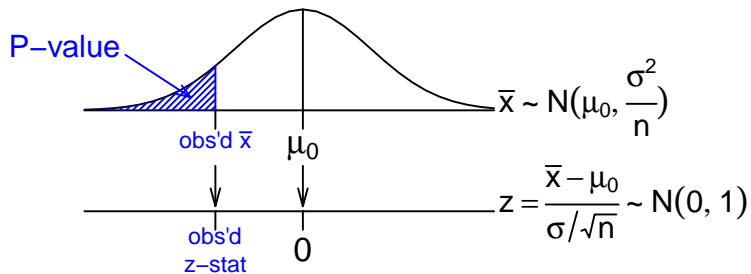
$$z\text{-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} < -z_\alpha.$$

P-values for Lower One-Sided Hypothesis Tests

To test $H_0: \mu = \mu_0$ v.s. **lower one-sided alternative** $H_1: \mu < \mu_0$, the P -value is the lower tail probability

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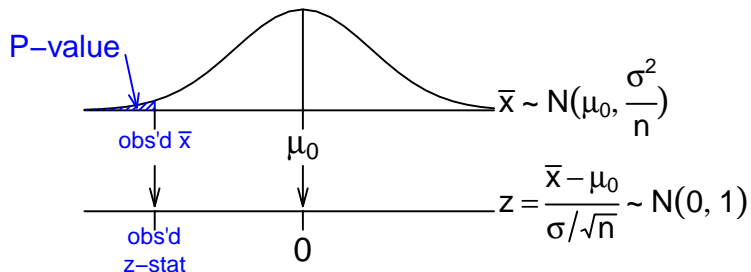


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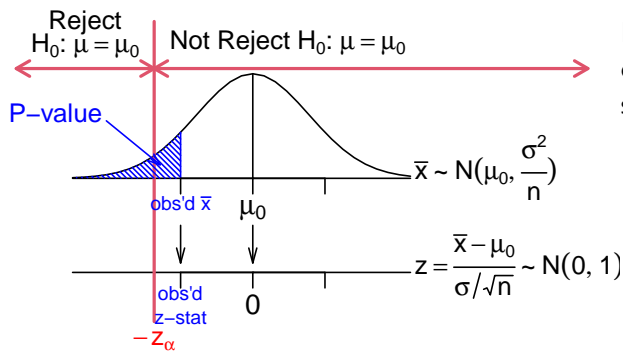
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P-value v.s. Critical Value for Lower One-Sided Tests

Observed that

if $z\text{-statistic} > -z_\alpha$ then $P\text{-value} > \alpha$



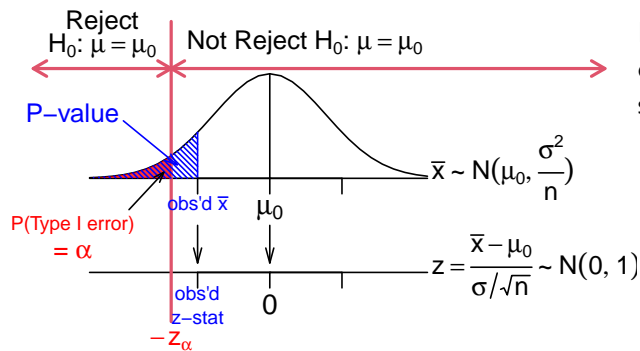
Two equivalent approaches to test $H_0: \mu = \mu_0$ v.s. $H_1: \mu < \mu_0$ control P(Type I error) at the significance level α :

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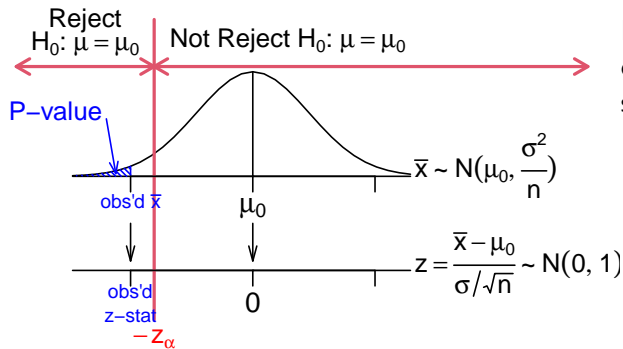
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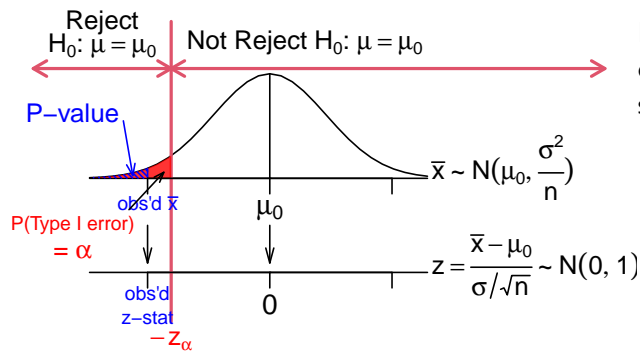
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P-value Approach or Critical Value Approach?

We introduced both the critical value approach and the *P*-value approach for hypothesis testing. They are equivalent but we generally *recommend the *P*-value approach*, for two reasons.

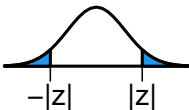

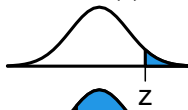
- ▶ The rejection rule is simpler, just compare the *P*-value with the significance level α
- ▶ More importantly, we can simply report the *P*-value and let people choose their own significance level $\alpha = P(\text{Type I error})$ and decide whether to reject or not to reject the H_0

Recap: 1- & 2-Sided Rejection Regions & P -values

For $H_0: \mu = \mu_0$, $z\text{-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1)$, reject H_0 at level α if

- ▶ $z\text{-stat} > z_\alpha$ for $H_1: \mu > \mu_0$
- ▶ $z\text{-stat} < -z_\alpha$ for $H_1: \mu < \mu_0$
- ▶ $|z\text{-stat}| > z_{\alpha/2}$ for $H_1: \mu \neq \mu_0$

The P -values are as follows, where the bell-shape curve is the standard normal curve

H_1	two-sided $\mu \neq \mu_0$	lower one-sided $\mu < \mu_0$	upper one-sided $\mu > \mu_0$
$P\text{-value}$	$2(1 - \Phi(z))$ 	$\Phi(z)$ 	$1 - \Phi(z)$ 

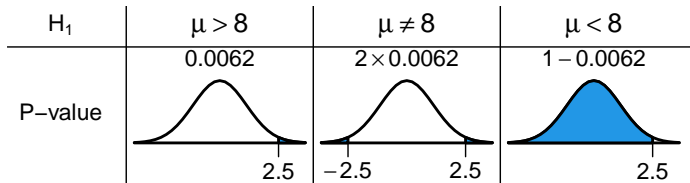
Example w/ Data

Data: $X_1, \dots, X_{100} \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2 = 6^2)$, w/ sample mean $\bar{x} = 9.5$.

For $H_0: \mu = 8$,

$$z\text{-stat} = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} = \frac{9.5 - 8}{6/\sqrt{100}} = \frac{1.5}{0.6} = 2.5,$$

$$P\text{-value} = \begin{cases} 1 - \Phi(2.5) \approx 0.0062 & \text{if } H_1: \mu > 8 \\ 2(1 - \Phi(2.5)) \approx 0.0124 & \text{if } H_1: \mu \neq 8 \\ \Phi(2.5) \approx 1 - 0.0062 = 0.9938 & \text{if } H_1: \mu < 8 \end{cases}$$



For $H_1: \mu > 8$ or $\mu \neq 8$, we reject H_0 since $P\text{-value} < 5\%$.

For $H_1: \mu < 8$, no reason to reject $H_0: \mu = 8$ since $H_1: \mu < 8$ is less plausible than

One Sample Tests for Mean, Unknown σ^2

One Sample Tests for Mean (Unknown σ^2) — Rejection Regions

Data: $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

The test statistic for testing $H_0: \mu = \mu_0$ with unknown σ^2 is

$$T = \frac{\bar{X} - \mu_0}{\sqrt{S^2/n}}, \quad \text{where} \quad S^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}.$$

Under $H_0: \mu = \mu_0$, $T \sim t_{n-1}$, we reject H_0 at level α if

- ▶ $t\text{-stat} > t_{n-1, \alpha}$ for $H_1: \mu > \mu_0$
- ▶ $t\text{-stat} < -t_{n-1, \alpha}$ for $H_1: \mu < \mu_0$
- ▶ $|t\text{-stat}| > t_{n-1, \alpha/2}$ for $H_1: \mu \neq \mu_0$

where $t\text{-stat}$ is the observed value of T

$$t\text{-stat} = \frac{\bar{x} - \mu_0}{\sqrt{s^2/n}}, \quad \text{in which} \quad s^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}.$$

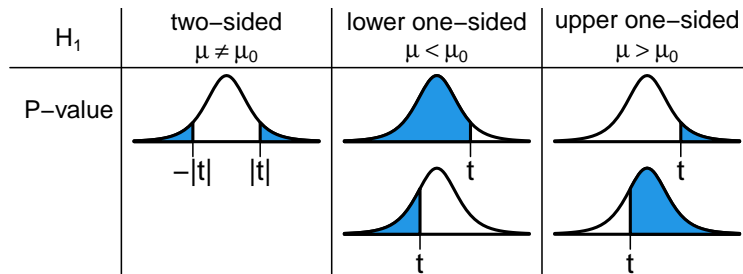
and $t_{n-1, \alpha}$ satisfies

$$P(T > t_{n-1, \alpha}) = \alpha \quad \text{for } T \sim t_{n-1}.$$

One Sample Tests for Mean (Unknown σ^2) — P -values

The P -values for testing $H_0: \mu = \mu_0$ with unknown σ^2 is

$$P\text{-value} = \begin{cases} P(T > t\text{-stat}) & \text{if } H_1: \mu > \mu_0 \\ P(|T| > |t\text{-stat}|) = 2P(T > |t\text{-stat}|) & \text{if } H_1: \mu \neq \mu_0 \\ P(T < t\text{-stat}) & \text{if } H_1: \mu < \mu_0 \end{cases}$$



The bell-shape curve above is the t -curve with $df = n - 1$, not the normal curve. We reject H_0 when $P\text{-value} < \alpha$.

One Sample Test for Variance

One Sample Test for Variance — Test Statistic

Data: $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

The test statistic for testing $H_0: \sigma^2 = \sigma_0^2$ with unknown μ is

$$V = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma_0^2} = \frac{(n-1)S^2}{\sigma_0^2}.$$

- ▶ What's the distribution of V under $H_0: \sigma^2 = \sigma_0^2$?
- ▶ $V \geq 0$
- ▶ Large V far above 1 is evidence for $H_1: \sigma^2 > \sigma_0^2$
- ▶ V far below 1 is evidence for $H_1: \sigma^2 < \sigma_0^2$
- ▶ V being far from 1 is evidence for $H_1: \sigma^2 \neq \sigma_0^2$

One Sample Test for Variance — Test Statistic

Data: $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

The test statistic for testing $H_0: \sigma^2 = \sigma_0^2$ with unknown μ is

$$V = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma_0^2} = \frac{(n-1)S^2}{\sigma_0^2}.$$

► What's the distribution of V under $H_0: \sigma^2 = \sigma_0^2$?

$V \sim \chi_{n-1}^2$, a **chi-squared** distribution w/ $n-1$ degrees of freedom

► $V \geq 0$

► Large V far above 1 is evidence for $H_1: \sigma^2 > \sigma_0^2$

► V far below 1 is evidence for $H_1: \sigma^2 < \sigma_0^2$

► V being far from 1 is evidence for $H_1: \sigma^2 \neq \sigma_0^2$

One Sample Test of Equal Variance — Rejection Region

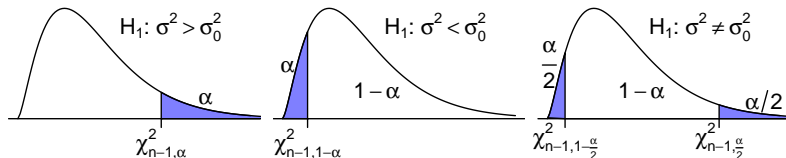
We reject H_0 at level α if

- ▶ $v\text{-stat} > \chi_{n-1,\alpha}^2$ for $H_1: \sigma^2 > \sigma_0^2$
- ▶ $v\text{-stat} < \chi_{n-1,1-\alpha}^2$ for $H_1: \sigma^2 < \sigma_0^2$
- ▶ $v\text{-stat} > \chi_{n-1,\alpha/2}^2$ or $v\text{-stat} < \chi_{n-1,1-\alpha/2}^2$ or for $H_1: \sigma^2 \neq \sigma_0^2$

where $v\text{-stat}$ is the observed value of V

$$v\text{-stat} = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{\sigma_0^2}.$$

and $\chi_{n-1,\alpha}^2$ satisfies $P(V > \chi_{n-1,\alpha}^2) = \alpha$ for $V \sim \chi_{n-1}^2$.

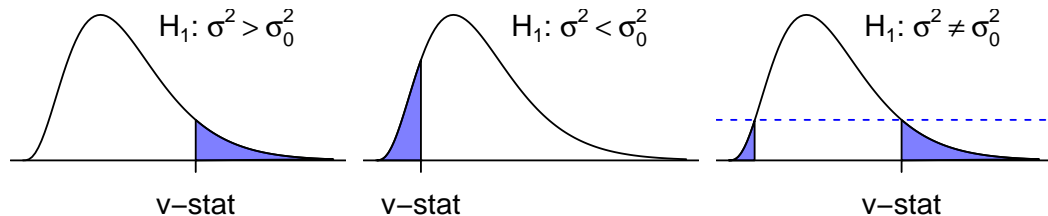


One Sample Test for Variance — P -value

The P -values for testing $H_0: \sigma^2 = \sigma_0^2$ with unknown μ is

$$P\text{-value} = \begin{cases} P(V > v\text{-stat}) & \text{if } H_1: \sigma^2 > \sigma_0^2 \\ P(V < v\text{-stat}) & \text{if } H_1: \sigma^2 < \sigma_0^2 \end{cases}$$

What's the two-sided P -value?



Two Sample Tests for Mean (Equal Variance)

Two Sample Test for Mean (Equal Variance) — Test Statistic

Consider two normal random samples of size n_1 and n_2 respectively

$$\left. \begin{array}{l} X_{11}, X_{12}, \dots, X_{1n_1} \\ X_{21}, X_{22}, \dots, X_{2n_2} \end{array} \right\} \begin{array}{l} \stackrel{\text{iid}}{\sim} N(\mu_1, \sigma^2) \\ \stackrel{\text{iid}}{\sim} N(\mu_2, \sigma^2) \end{array} \rightarrow \text{indep., same } \sigma^2.$$

For testing $H_0: \mu_1 = \mu_2$, the two-sample T-statistic is

$$T = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{(\frac{1}{n_1} + \frac{1}{n_2})S^2}}, \text{ where } S^2 = \frac{\sum_{i=1}^{n_1} (X_{1i} - \bar{X}_1)^2 + \sum_{j=1}^{n_2} (X_{2j} - \bar{X}_2)^2}{n_1 + n_2 - 2}$$

Under $H_0: \mu_1 = \mu_2$, $T \sim t_{n_1+n_2-2}$.

Two Sample Test for Mean (Equal Variance) — Rejection Region

We reject $H_0: \mu_1 = \mu_2$ at level α if

- ▶ $t\text{-stat} > t_{n_1+n_2-2, \alpha}$ for $H_1: \mu_1 > \mu_2$
- ▶ $t\text{-stat} < -t_{n_1+n_2-2, \alpha}$ for $H_1: \mu_1 < \mu_2$
- ▶ $|t\text{-stat}| > t_{n_1+n_2-2, \alpha/2}$ for $H_1: \mu_1 \neq \mu_2$

where $t\text{-stat}$ is the observed value of T

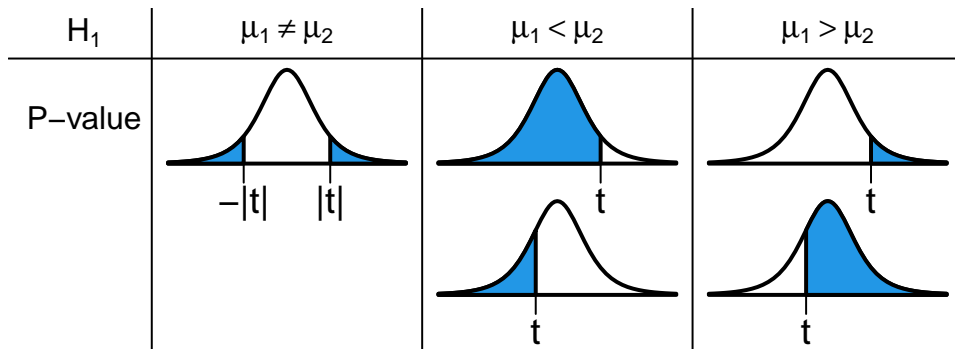
$$t\text{-stat} = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{(\frac{1}{n_1} + \frac{1}{n_2})s^2}}, \text{ in which } s^2 = \frac{\sum_{i=1}^{n_1} (x_{1i} - \bar{x}_1)^2 + \sum_{j=1}^{n_2} (x_{2j} - \bar{x}_2)^2}{n_1 + n_2 - 2}.$$

and $t_{n_1+n_2-2, \alpha}$ satisfies

$$P(T > t_{n_1+n_2-2, \alpha}) = \alpha \quad \text{for } T \sim t_{n_1+n_2-2}.$$

In L17, we show that a two-sided two-sample test for mean is equivalent to the GLR test.

Two Sample Test for Mean (Equal Variance) — P -Value



The bell curve above is the t -curve with $n_1 + n_2 - 2$ degrees of freedom.

Two Sample Tests for Mean (Unequal Variance)

Two Sample Test for Mean (Unequal Variance)

Without the equal variance assumption, by the indep of the two samples, we know

$$\text{Var}(\bar{X}_1 - \bar{X}_2) = \text{Var}(\bar{X}_1) + \text{Var}(\bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

The two-sample T -statistic without the equal variance assumption is

$$T = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}} \quad \text{where} \quad \begin{aligned} S_1^2 &= \frac{\sum_{i=1}^{n_1} (X_{1i} - \bar{X}_1)^2}{n_1 - 1} \\ S_2^2 &= \frac{\sum_{j=1}^{n_2} (X_{2j} - \bar{X}_2)^2}{n_2 - 1} \end{aligned}$$

- ▶ Unfortunately, the T -statistic above does NOT have a t -distribution, even under $H_0: \mu_1 - \mu_2$
- ▶ Fortunately, it can be approximated by a t -distribution with a certain degrees of freedom.

See the next slide for the approximation

Approximate Distribution of the Two-Sample t -Statistic

Under $H_0: \mu_1 - \mu_2$, the two-sample t -statistic has an **approximate t_k distribution**, with the degrees of freedom k as follows

$$k = \frac{(w_1 + w_2)^2}{w_1^2/(n_1 - 1) + w_2^2/(n_2 - 1)}, \quad \text{where} \quad \begin{aligned} w_1 &= s_1^2/n_1, \\ w_2 &= s_2^2/n_2. \end{aligned}$$

The rejection regions and the calculation of the P -value are similar to the equal variance case, except for the degrees of freedom and thus is not repeated here.

Two Sample Tests of Equal Variance

Two Sample Tests of Equal Variance

Consider two normal random samples of size n_1 and n_2 respectively

$$\left. \begin{array}{l} X_{11}, X_{12}, \dots, X_{1n_1} \\ X_{21}, X_{22}, \dots, X_{2n_2} \end{array} \right\} \begin{array}{l} \stackrel{\text{iid}}{\sim} N(\mu_1, \sigma_1^2) \\ \stackrel{\text{iid}}{\sim} N(\mu_2, \sigma_2^2) \end{array} \rightarrow \text{indep.}$$

For testing $H_0: \sigma_1^2 = \sigma_2^2$, the test-statistic is

$$F = \frac{S_1^2}{S_2^2} \quad \text{where } S_k^2 = \frac{\sum_{i=1}^{n_k} (X_{ki} - \bar{X}_k)^2}{n_k - 1}, \quad k = 1, 2.$$

- ▶ What's the distribution of F under $H_0: \sigma_1^2 = \sigma_2^2$?
 - ▶ $(n_1 - 1)S_1^2/\sigma_1^2 \sim \chi_{n_1-1}^2$ and $(n_2 - 1)S_2^2/\sigma_2^2 \sim \chi_{n_2-1}^2$ are indep
 - ▶ So $F \sim F_{n_1-1, n_2-1}$ has an F -distribution w/ $n_1 - 1$ and $n_2 - 1$ degrees of freedom under $H_0: \sigma_1^2 = \sigma_2^2$
- ▶ $F \geq 0$
- ▶ F far above 1 is evidence for $H_1: \sigma_1^2 > \sigma_2^2$
- ▶ F far below 1 is evidence for $H_1: \sigma_1^2 < \sigma_2^2$
- ▶ F being far away from 1 is evidence for $H_1: \sigma^2 \neq \sigma_0^2$

Two-Sample Test for Variance — Rejection Region

We reject H_0 at level α if

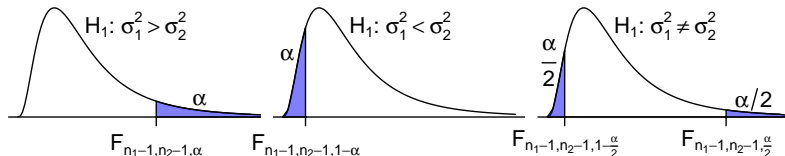
- ▶ $f\text{-stat} > F_{n_1-1, n_2-2, \alpha}$ for $H_1: \sigma_1^2 > \sigma_2^2$
- ▶ $f\text{-stat} < F_{n_1-1, n_2-2, 1-\alpha}$ for $H_1: \sigma_1^2 < \sigma_2^2$
- ▶ $f\text{-stat} > F_{n_1-1, n_2-2, \alpha/2}$ or $f\text{-stat} < F_{n_1-1, n_2-1, 1-\alpha/2}$ for $H_1: \sigma_1^2 \neq \sigma_2^2$

where $f\text{-stat}$ is the observed value of V

$$f\text{-stat} = \frac{s_1^2}{s_2^2}.$$

and $F_{n_1-1, n_2-1, \alpha}$ satisfies

$$P(F > F_{n_1-1, n_2-1, \alpha}) = \alpha \quad \text{for } F \sim F_{n_1-1, n_2-1}.$$

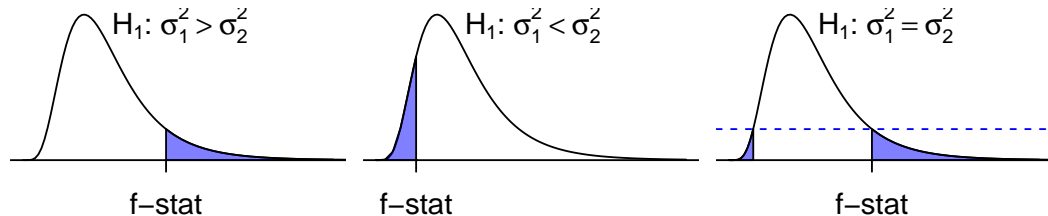


Two-Sample Test for Equal Variance — P -value

The P -values for testing $H_0: \sigma_1^2 = \sigma_2^2$ with unknown μ is

$$P\text{-value} = \begin{cases} P(F > f\text{-stat}) & \text{if } H_1: \sigma_1^2 > \sigma_2^2 \\ P(F < f\text{-stat}) & \text{if } H_1: \sigma_1^2 < \sigma_2^2 \end{cases}$$

What's the two-sided P -value?



Robustness to Non-Normality

If the data are **NOT normally** distributed,

- ▶ the 4 tests about the mean are approx. valid if the sample size is large by CLT.
- ▶ the 2 tests about the variance are no longer valid