

Stat 246 Homework V - Solution

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1 [60 points]

(a)

The joint distribution of X_0, X_1, X_2 has the following form.

$$P(x_0, x_1, x_2) = \frac{1}{Z} \exp\left(\sum_{c \in \mathcal{C}} \psi_c(x_c)\right),$$

where $\mathcal{C} = \{(0, 1), (1, 2)\}$, the set of all maximal cliques. The normalizing constant Z is

$$\begin{aligned} Z &= \sum_{x_0} \sum_{x_1} \sum_{x_2} \exp(\alpha x_0 x_1 + \alpha x_1 x_2) \\ &= \sum_{x_0} \sum_{x_2} \{1 + \exp(\alpha x_0 + \alpha x_2)\} \\ &= \sum_{x_0} \{2 + \exp(\alpha x_0) + \exp(\alpha x_0 + \alpha)\} \\ &= 5 + 2 \exp(\alpha) + \exp(2\alpha). \end{aligned}$$

The marginal distribution on X_0 is

$$\begin{aligned} P(x_0) &= \sum_{x_1} \sum_{x_2} P(x_0, x_1, x_2) \\ &= \sum_{x_1} \sum_{x_2} \exp(\alpha x_0 x_1 + \alpha x_1 x_2) / Z \\ &= \frac{2 + \exp(\alpha x_0) + \exp(\alpha(x_0 + 1))}{5 + 2 \exp(\alpha) + \exp(2\alpha)}. \end{aligned}$$

(b)

We have the joint distribution of X_0, X_1, X_2 given by

$$\begin{aligned} P(x_0, x_1, x_2) &= \frac{\exp(\alpha x_0 x_1 + \alpha x_1 x_2)}{5 + 2 \exp(\alpha) + \exp(2\alpha)} \\ &= \frac{\exp(\alpha(x_0 x_1 + x_1 x_2 - 2))}{5 \exp(-2\alpha) + 2 \exp(-\alpha) + 1}. \end{aligned}$$

Therefore,

$$P(x_0, x_1, x_2) \rightarrow \frac{1}{8} \quad \text{as } \alpha \rightarrow 0,$$

$$P(x_0, x_1, x_2) \rightarrow \begin{cases} 1, & \text{if } x_0 = x_1 = x_2 = 1, \\ 0, & \text{otherwise,} \end{cases} \quad \text{as } \alpha \rightarrow \infty.$$

(c)

Suppose the rows and columns of $M = \begin{pmatrix} 1 & 1 \\ 1 & e^\alpha \end{pmatrix}$ are indexed by 0, 1, that is, $M_{0,0} = M_{0,1} = M_{1,0} = 1$ and $M_{1,1} = e^\alpha$. Then we have $M_{X_i, X_{i+1}} = \exp(\alpha X_i X_{i+1})$. Since $(AB)_{i,j} = \sum_k A_{i,k} B_{k,j}$ for any A, B , we have

$$\begin{aligned} Z &= \sum_{x_0, x_1, x_2} \exp(\alpha x_0 x_1 + \alpha x_1 x_2) \\ &= \sum_{x_0, x_2} \sum_{x_1} M_{x_0, x_1} M_{x_1, x_2} \\ &= \sum_{x_0, x_2} M_{x_0, x_2}^2 \end{aligned}$$

which is the sum of all entries of M^2 .

(d)

For the graph G over n variables X_0, \dots, X_{n-1} with edges only between consecutive nodes, the joint distribution of X_0, \dots, X_{n-1} is

$$p(x_0, \dots, x_{n-1}) = \frac{1}{Z} \exp(\alpha x_0 x_1 + \dots + \alpha x_{n-2} x_{n-1}) = \frac{1}{Z} \prod_{k=0}^{n-2} \exp(\alpha x_k x_{k+1}) = \frac{1}{Z} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}}.$$

Then the normalizing constant Z is

$$\begin{aligned} Z &= \sum_{x_0, \dots, x_{n-1}} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}} \\ &= \sum_{x_0, x_{n-1}} \sum_{x_1, \dots, x_{n-2}} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}} \\ &= \sum_{x_0, x_{n-1}} M_{x_0, x_{n-1}}^{n-1}, \end{aligned}$$

since

$$\begin{aligned}
\sum_{x_1, \dots, x_{n-2}} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}} &= \sum_{x_2, \dots, x_{n-2}} \prod_{k=2}^{n-2} M_{x_k, x_{k+1}} \underbrace{\sum_{x_1} M_{x_0, x_1} M_{x_1, x_2}}_{M_{x_0, x_2}^2} \\
&= \sum_{x_3, \dots, x_{n-2}} \prod_{k=3}^{n-2} M_{x_k, x_{k+1}} \underbrace{\sum_{x_2} M_{x_0, x_2}^2 M_{x_2, x_3}}_{M_{x_0, x_3}^3} \\
&\quad \vdots \\
&= M_{x_0, x_{n-1}}^{n-1}.
\end{aligned}$$

Therefore, the normalizing constant Z is the sum of all entries of M^{n-1} .

The marginal distribution on X_0 is obtained as follows.

$$\begin{aligned}
P(x_0) &= \sum_{x_1, \dots, x_{n-1}} P(x_0, \dots, x_{n-1}) \\
&= \sum_{x_1, \dots, x_{n-1}} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}} / Z \\
&= \sum_{x_{n-1}} \sum_{x_1, \dots, x_{n-2}} \prod_{k=0}^{n-2} M_{x_k, x_{k+1}} / Z \\
&= \sum_{x_{n-1}} M_{x_0, x_{n-1}}^{n-1} / Z \\
&= (M_{x_0, 0}^{n-1} + M_{x_0, 1}^{n-1}) / \sum_{y, y'=0,1} M_{y, y'}^{n-1}.
\end{aligned}$$

(e)

For any joint distribution $p(x_1, \dots, x_n)$ on n variables,

$$\begin{aligned}
p(x_1, \dots, x_n) &= \frac{p(x_1, \dots, x_n)}{p(x_1, \dots, x_{n-1})} \frac{p(x_1, \dots, x_{n-1})}{p(x_1, \dots, x_{n-2})} \dots \frac{p(x_1, x_2)}{p(x_1)} \\
&= p(x_n | x_1, \dots, x_{n-1}) p(x_{n-1} | x_1, \dots, x_{n-2}) \dots p(x_2 | x_1) p(x_1).
\end{aligned}$$

Suppose the joint distribution is given as $p(x_1, \dots, x_n) = \frac{1}{Z} \prod_{k=1}^{n-1} M_{x_k, x_{k+1}}$ as in (d). Then

$$\begin{aligned}
p(x_n | x_1, \dots, x_{n-1}) &= \frac{p(x_1, \dots, x_{n-1}, x_n)}{p(x_1, \dots, x_{n-1})} \\
&= \frac{\prod_{k=1}^{n-1} M_{x_k, x_{k+1}} / Z}{\sum_{x_n} \prod_{k=1}^{n-1} M_{x_k, x_{k+1}} / Z} \\
&= \frac{M_{x_{n-1}, x_n} \prod_{k=1}^{n-2} M_{x_k, x_{k+1}}}{\sum_{x_n} M_{x_{n-1}, x_n} \prod_{k=1}^{n-2} M_{x_k, x_{k+1}}} \\
&= \frac{M_{x_{n-1}, x_n}}{\sum_{x_n} M_{x_{n-1}, x_n}} \\
&= \frac{\exp(\alpha x_{n-1} x_n)}{1 + \exp(\alpha x_{n-1})}.
\end{aligned}$$

Since this is a function of x_{n-1} and x_n only,

$$p(x_n | x_1, \dots, x_{n-1}) = p(x_n | x_{n-1}).$$

Similarly for $i = 2, \dots, n-1$,

$$\begin{aligned}
p(x_i | x_1, \dots, x_{i-1}) &= \frac{p(x_1, \dots, x_{i-1}, x_i)}{p(x_1, \dots, x_{i-1})} \\
&= \frac{\sum_{x_{i+1}, \dots, x_n} \prod_{k=1}^{n-1} M_{x_k, x_{k+1}} / Z}{\sum_{x_i, x_{i+1}, \dots, x_n} \prod_{k=1}^{n-1} M_{x_k, x_{k+1}} / Z} \\
&= \frac{\prod_{k=1}^{i-1} M_{x_k, x_{k+1}} \sum_{x_{i+1}, \dots, x_n} \prod_{k=i}^{n-1} M_{x_k, x_{k+1}}}{\prod_{k=1}^{i-2} M_{x_k, x_{k+1}} \sum_{x_i, x_{i+1}, \dots, x_n} \prod_{k=i-1}^{n-1} M_{x_k, x_{k+1}}} \\
&= \frac{M_{x_{i-1}, x_i} \sum_{x_{i+1}, \dots, x_n} \prod_{k=i}^{n-1} M_{x_k, x_{k+1}}}{\sum_{x_i} M_{x_{i-1}, x_i} \sum_{x_{i+1}, \dots, x_n} \prod_{k=i}^{n-1} M_{x_k, x_{k+1}}} \\
&= \frac{M_{x_{i-1}, x_i} \sum_{x_n} M_{x_i, x_n}^{n-i}}{\sum_{x_i} M_{x_{i-1}, x_i} \sum_{x_n} M_{x_i, x_n}^{n-i}}, \quad (\text{since } \sum_{x_{i+1}, \dots, x_{n-1}} \prod_{k=i}^{n-1} M_{x_k, x_{k+1}} = M_{x_i, x_n}^{n-i})
\end{aligned}$$

which is a function of x_{i-1} and x_i only, so we have

$$p(x_i | x_1, \dots, x_{i-1}) = p(x_i | x_{i-1}),$$

and

$$p(x_1, \dots, x_n) = p(x_n | x_{n-1}) p(x_{n-1} | x_{n-2}) \dots p(x_2 | x_1) p(x_1).$$

Thus, this joint distribution defines a Markov Chain. However, since each transition probability $p(x_i | x_{i-1}) := P(X_i = x_i | X_{i-1} = x_{i-1})$ depends on M^{n-i} which is different for each i , the transition probabilities are not the same at each step, so the Markov Chain is not homogeneous.

Note: We can also use the conditional independence properties defined by the graph to show that the joint distribution defines a Markov Chain. Since the joint distribution is defined as the product of clique functions and the normalizing constant, it has a Markov property with respect to the graph, that is, if C separates A from B , then X_A and X_B are conditionally independent given X_C . Since node $i-1$ separates node i from nodes $1, \dots, i-2$, X_i is conditionally independent of X_1, \dots, X_{i-2} given X_{i-1} , so $p(x_i | x_1, \dots, x_{i-1}) = p(x_i | x_{i-1})$. However, you still need to calculate the transition probabilities to see if the Markov Chain is homogeneous.

(f)

Without clique functions for singletons, we have

$$\begin{aligned} P(x_0, x_1, x_2) &= \exp(\alpha x_0 x_1 + \alpha x_1 x_2) / Z_0 \\ &= \exp(-x^t \Gamma_0 x) / Z_0, \end{aligned}$$

where $x = (x_0, x_1, x_2)^t$ and $\Gamma_0 = \begin{pmatrix} 0 & -\alpha/2 & 0 \\ -\alpha/2 & 0 & -\alpha/2 \\ 0 & -\alpha/2 & 0 \end{pmatrix}$. Since Γ_0 is not a positive definite matrix, we cannot compute Z_0 .

If we add a clique function for each singleton as given, then

$$\begin{aligned} P(x_0, x_1, x_2) &= \exp(\alpha x_0 x_1 + \alpha x_1 x_2 - \beta x_0^2 - \beta x_1^2 - \beta x_2^2) / Z \\ &= \exp(-x^t \Gamma x) / Z, \end{aligned}$$

where $\Gamma = \begin{pmatrix} \beta & -\alpha/2 & 0 \\ -\alpha/2 & \beta & -\alpha/2 \\ 0 & -\alpha/2 & \beta \end{pmatrix}$. Note that the eigenvalues of Γ are $\beta, \beta \pm \alpha/\sqrt{2}$. Thus, to make Γ positive definite, we need to assume that $\beta - \alpha/\sqrt{2} > 0$, that is, $\beta^2 > \alpha^2/2$. Then this joint distribution is a multivariate normal distribution with mean $\mu = (0, 0, 0)$ and the covariance matrix $\Sigma = \Gamma^{-1}/2$, so the normalizing constant is

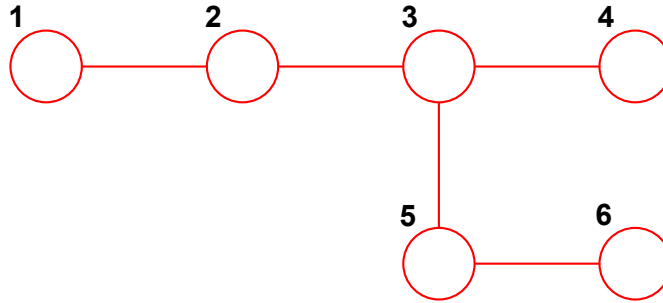
$$Z = |2\pi\Sigma|^{1/2} = |2\pi\Gamma^{-1}/2|^{1/2} = \pi^{3/2} |\Gamma|^{-1/2} = \frac{\pi \sqrt{\pi}}{\sqrt{\beta(\beta + \alpha/\sqrt{2})(\beta - \alpha/\sqrt{2})}} = \frac{\pi \sqrt{\pi}}{\sqrt{\beta(\beta^2 - \alpha^2/2)}}.$$

The joint distribution is then,

$$P(x_0, x_1, x_2) = \frac{\sqrt{\beta(\beta^2 - \alpha^2/2)}}{\pi \sqrt{\pi}} \exp(\alpha x_0 x_1 + \alpha x_1 x_2 - \beta x_0^2 - \beta x_1^2 - \beta x_2^2).$$

2 [40 points]

(a)



(b)

Since the maximal cliques are $\mathcal{C} = \{(1, 2), (2, 3), (3, 4), (3, 5), (5, 6)\}$, the joint distribution is

$$\begin{aligned} p(x) &= \frac{1}{Z} \exp \left\{ \sum_{c \in \mathcal{C}} \psi_c(x_c) \right\} \\ &= \frac{1}{Z} \exp \{ \psi_{12}(x_1, x_2) + \psi_{23}(x_2, x_3) + \psi_{34}(x_3, x_4) + \psi_{35}(x_3, x_5) + \psi_{56}(x_5, x_6) \}. \end{aligned}$$

(c)

The normalizing constant is

$$Z = \sum_{x_1} \dots \sum_{x_6} \exp \{ \psi_{12}(x_1, x_2) + \psi_{23}(x_2, x_3) + \psi_{34}(x_3, x_4) + \psi_{35}(x_3, x_5) + \psi_{56}(x_5, x_6) \}.$$

$\exp \{ \psi_{12}(x_1, x_2) + \dots + \psi_{56}(x_5, x_6) \}$ can be represented as a set of numbers, one for each possible value for $x = (x_1, \dots, x_6)$. Since there are 6 variables each with K states, there are K^6 values for x , so the summation over all sequences x takes $K^6 - 1 = O(K^6)$ operations.

(d)

Let $f(\bullet) = \exp\{\psi(\bullet)\}$. We can rearrange the order of the summations and the multiplications as follows.

$$\begin{aligned}
Z &= \sum_{x_1} \dots \sum_{x_6} \exp\{\psi_{12}(x_1, x_2) + \psi_{23}(x_2, x_3) + \psi_{34}(x_3, x_4) + \psi_{35}(x_3, x_5) + \psi_{45}(x_4, x_5)\} \\
&= \sum_{x_1} \dots \sum_{x_6} f_{12}(x_1, x_2) f_{23}(x_2, x_3) f_{34}(x_3, x_4) f_{35}(x_3, x_5) f_{56}(x_5, x_6) \\
&= \sum_{x_1} \sum_{x_2} f_{12}(x_1, x_2) \underbrace{\sum_{x_3} f_{23}(x_2, x_3)}_{O(K^2)} \underbrace{\sum_{x_4} f_{34}(x_3, x_4)}_{O(K^2)} \underbrace{\sum_{x_5} f_{35}(x_3, x_5)}_{O(K^2)} \underbrace{\sum_{x_6} f_{56}(x_5, x_6)}_{O(K^2)} \\
&\quad \underbrace{\hspace{10em}}_{O(K^2)} \underbrace{\hspace{10em}}_{O(K^2)} \\
&\quad \underbrace{\hspace{10em}}_{O(K^2)} \\
&\quad \underbrace{\hspace{10em}}_{O(K^2)} \\
&\quad \underbrace{\hspace{10em}}_{O(K^2)} \\
&\quad \underbrace{\hspace{10em}}_{O(K)}
\end{aligned}$$

Note that you can start with $\sum_{x_6} f_{56}(x_5, x_6)$ or $\sum_{x_4} f_{34}(x_3, x_4)$ or $\sum_{x_1} f_{12}(x_1, x_2)$, i.e. start by removing a node that has only one neighbor.

Using the above reordered expression, the computational cost is as follows.

The summation over x_6 involves only the function $f_{56}(x_5, x_6)$, which is a table of $K \times K$ numbers. We have to sum this table over x_6 for each values of x_5 and so this summation has $K(K - 1)$ cost.

The resulting vector of K numbers $\sum_{x_6} f_{56}(x_5, x_6)$, $x_5 = 1, \dots, K$, is multiplied by the matrix of numbers $f_{35}(x_3, x_5)$ and so this has K^2 cost.

Since there are 5 summations and 4 multiplications of these kinds and the last summation over x_1 takes only $K - 1$ operations, the total cost of evaluating Z is about $9K^2$.

Note: For a chain of N nodes, evaluating the normalizing constant has $O(NK^2)$ cost which is linear in the length of the chain. See p.396. Although the graph given here is not a chain, the calculation is done in a similar way. As long as the graph maintains a tree shape as it grows to N nodes, the evaluation of the normalizing constant takes $O(NK^2)$ operations.