

# Sparse correlation screening in high dimension

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# Acknowledgements

- Bala Rajaratnam (Stanford)
- Isaac Newton Institute
- DIGITEO, Paris France
- NSF: ITR CCR-032557

1 Motivation

2 Theory

3 Application

4 Conclusions

# Outline

- 1 Motivation
- 2 Theory
- 3 Application
- 4 Conclusions

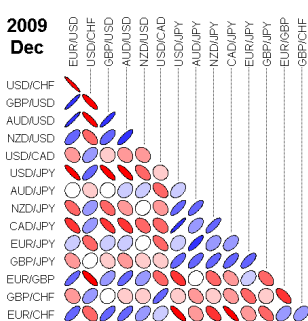
# Correlation analysis of financial time series



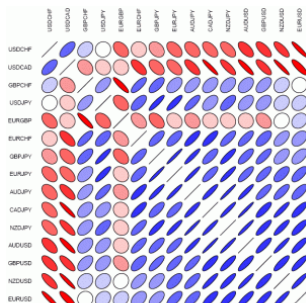
Source: FuturesMag.com

[www.futuresmag.com/.../Dom%20FEB%2024.JPG](http://www.futuresmag.com/.../Dom%20FEB%2024.JPG)

# $p$ -variate correlation analysis of financial data



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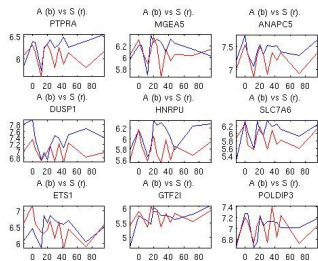
Sample covariance matrix:

$$\hat{\Sigma} = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{x}_i - \hat{\mu})(\mathbf{x}_i - \hat{\mu})^T$$

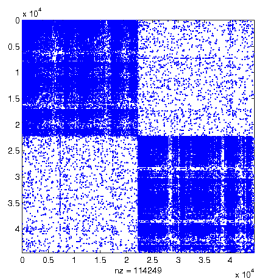
Sample correlation matrix:

$$\mathbf{R} = \mathbf{D}^{-1/2} \hat{\Sigma} \mathbf{D}^{-1/2}$$

# Correlation analysis of gene expression arrays



Gene expression profiles



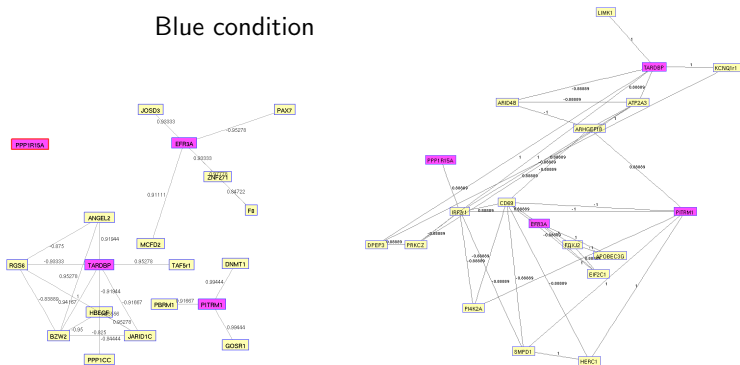
Correlation matrix  $R$





# Correlation screening and hub discovery

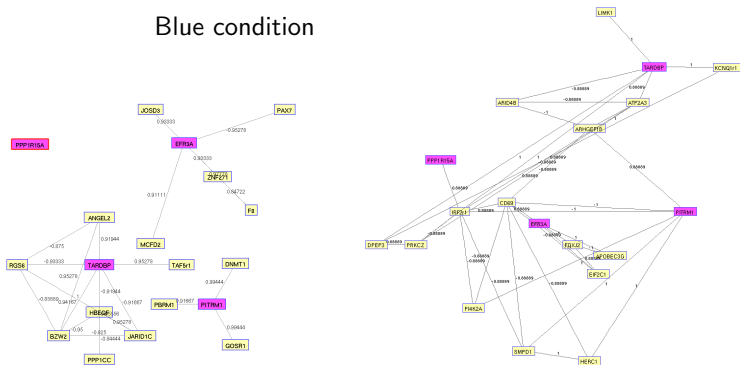
## Blue condition



- Correlation screening finds hubs of high sample correlation

# Correlation screening and hub discovery

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- Correlation screening finds hubs of high sample correlation
- Persistent correlation screening finds hubs surviving both treatments



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**Objective:** establish asymptotic (large  $p$ ) theory.

# Previous work

- Regularized  $l_2$  or  $l_{\mathcal{F}}$  covariance estimation
  - Shrinkage towards identity: Ledoit-Wolf (2005), Chen-Weisel-Eldar-H (2010)
  - Shrinkage towards banded: Bickel-Levina (2008)
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New framework: screening for highly correlated variables

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Scalable: computational complexity can be as low as  $O(\log p)$

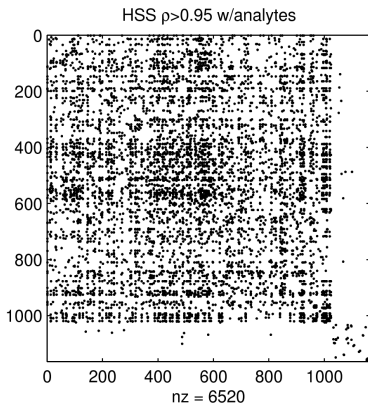
# Correlation screening = screening rows (variables) of $\mathbf{R}$

- For  $\mathbf{r}_{ij} = (\mathbf{R})_{ij}$  let  $\rho$  be a user-defined threshold in  $[0, 1]$
- Variable  $i$  passes correlation screen if:  $\max_{j \neq i} |\mathbf{r}_{ij}| \geq \rho$



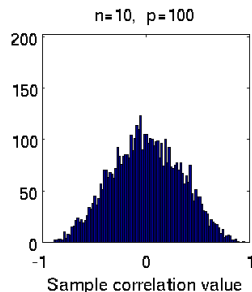
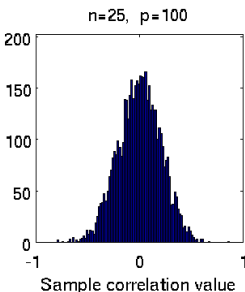
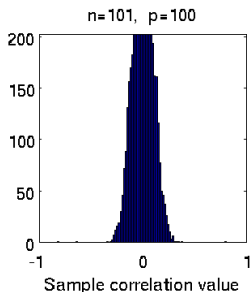
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- *Discovered* variables have high correlation with some other variable



# Phase transitions in correlation screening

- Number of discoveries exhibit phase transition phenomenon
- This phenomenon gets worse as  $p/n$  increases.



# Overview of mathematical results

Two types of results for auto-correlation and persistent correlation screening

- Characterize large  $p$  phase transition and its threshold.
- Poisson asymptotics for predicting false positive rates.

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- Poisson asymptotics for predicting false positive rates.

Main ingredients in our analysis

- Z-score representation:  $\mathbf{R} = \mathbf{U}^T \mathbf{U}$

$$\mathbf{U} = [\mathbf{U}_1, \dots, \mathbf{U}_p], \quad \mathbf{U}_i \in S_{n-2} \subset \mathbb{R}^{n-1}$$

- Geometric probability on unit-sphere  $S_{n-2}$
- Exchangeable process theory for dependent variables

# Sample correlation and Z-score distances

- Sample correlation between  $\mathbf{X}_i$  and  $\mathbf{X}_j$  is equal to Z-score inner product

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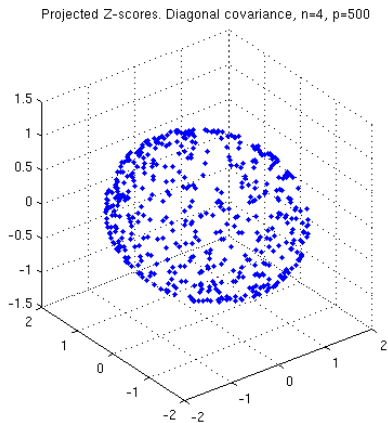
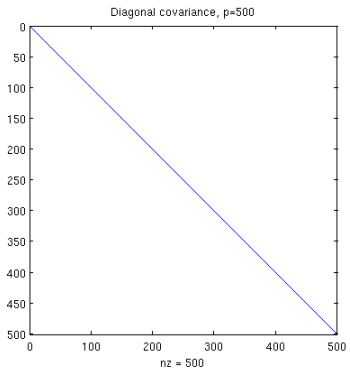
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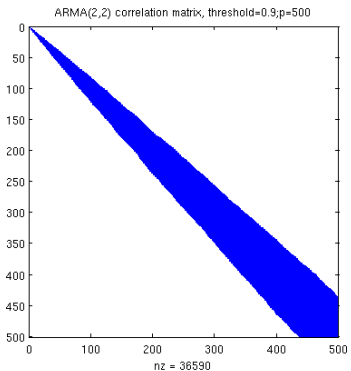
- Relate to Euclidean distance between  $\mathbf{U}_i$  and  $\mathbf{U}_j$

$$\|\mathbf{U}_i - \mathbf{U}_j\| = \sqrt{2(1 - \mathbf{r}_{ij})}$$

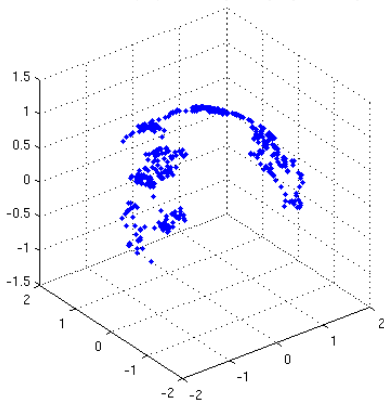
# Example: Z-scores for diagonal Gaussian



# Example : Z-scores for ARMA(2,2) Gaussian



Projected Z-scores. ARMA(2,2) model.  $a=[1.0,0.8], b=[1.0,-0.999], n=4,$





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# Mathematical analysis

Define  $N$  the number of discoveries:

$$N = \sum_{i=1}^p \phi_i$$

Where  $\phi = [\phi_1, \dots, \phi_p]$  is "discovery" indicator sequence:

$$\phi_i = \begin{cases} 1, & \max_{j \neq i} |\mathbf{r}_{ij}| \geq \rho \\ 0, & \text{o.w.} \end{cases}$$

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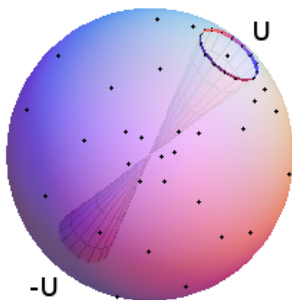
$$\phi_i = \begin{cases} 1, & \max_{j \neq i} |r_{ij}| \geq \rho \\ 0, & \text{o.w.} \end{cases}$$

**Objective:** Find mathematical expressions for  $E[N]$  as a function of  $p$ ,  $n$ ,  $\rho$ .

# Mathematical analysis

Conditional expectation of  $\phi_i$  has representation

$$E[\phi_i | \mathbf{U}_i] = P(\cup_{j \neq i} \mathbf{U}_j \in C_{\rho, \mathbf{U}_i} \cup C_{\rho, -\mathbf{U}_i} | \mathbf{U}_i)$$



# Mathematical analysis

Given  $\mathbf{U}_i$  define the binary sequence  $\mathbf{b} = [b_1, \dots, b_{p-1}]$

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Classical result [Thm. 4.5.4]{TW Anderson, 2003}:

## Lemma

*Let  $\mathbf{X}$  be a  $p$ -variate elliptical vector with diagonal dispersion matrix  $\Sigma$ . The Z-scores  $\{\mathbf{U}_i\}_{i=1}^p$  are i.i.d. random vectors uniformly distributed on  $S_{n-2}$ .*

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Implication:  $b_i$ 's are Bernoulli and  $E[\phi_i | \mathbf{U}_i] = 1 - (1 - P_0)^{p-1}$ .

# Main result: correlation screening

## Proposition

Let the  $n \times p$  data matrix  $\mathbb{X}$  have i.i.d. rows but possibly dependent columns. Let the sequence  $\{\rho_p\}_p$  of correlation thresholds be such that  $\rho_p \rightarrow 1$  and  $p(p-1)(1-\rho_p^2)^{(n-2)/2} \rightarrow d_n$  for some finite constant  $d_n$ . Then

$$\lim_{p \rightarrow \infty} E[N] = \kappa_n J(\overline{f_{\mathbf{u}_\bullet, \mathbf{u}_{*\bullet}}}), \quad (1)$$

where  $\kappa_n = a_n d_n / (n-2)$  and  $\overline{f_{\mathbf{u}_\bullet, \mathbf{u}_{*\bullet}}}$  is limit of average density

$$\overline{f_{\mathbf{u}_\bullet, \mathbf{u}_{*\bullet}}^{(p)}}(\mathbf{u}, \mathbf{v}) = \frac{1}{p} \sum_{i=1}^p \frac{1}{p-1} \sum_{j \neq i}^p \left( \frac{1}{2} f_{\mathbf{u}_i, \mathbf{u}_j}(\mathbf{u}, \mathbf{v}) + \frac{1}{2} f_{\mathbf{u}_i, \mathbf{u}_j}(\mathbf{u}, -\mathbf{v}) \right). \quad (2)$$



# Implication: uniform Z-score density is minimax

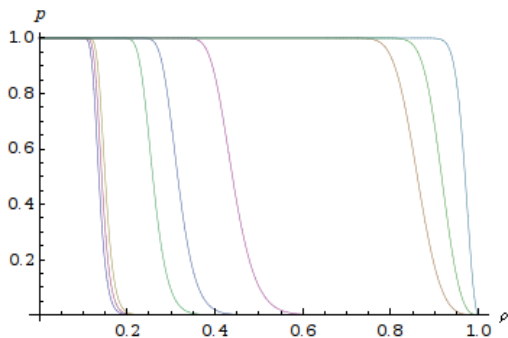
- $J(\overline{f_{\mathbf{U}, \mathbf{V}}})$ : related to Hellinger divergence and Rényi entropy

$$\begin{aligned}
 J(f_{\mathbf{U}, \mathbf{V}}) &= |S_{n-2}| \int f_{\mathbf{U}, \mathbf{V}}(\mathbf{w}, \mathbf{w}) d\mathbf{w} \\
 &= |S_{n-2}| \int (f_{\mathbf{U}|\mathbf{V}}(\mathbf{w}|\mathbf{w}) f_{\mathbf{V}|\mathbf{U}}(\mathbf{w}|\mathbf{w}))^{1/2} (f_{\mathbf{U}}(\mathbf{w}) f_{\mathbf{V}}(\mathbf{w}))^{1/2} d\mathbf{w} \\
 &\leq |S_{n-2}| \left( \int f_{\mathbf{U}|\mathbf{V}}(\mathbf{w}|\mathbf{w}) f_{\mathbf{V}|\mathbf{U}}(\mathbf{w}|\mathbf{w}) \right)^{1/2} \left( \int f_{\mathbf{U}}(\mathbf{w}) f_{\mathbf{V}}(\mathbf{w}) \right)^{1/2} \\
 &\leq H_2^{1/4}(f_{\mathbf{U}|\mathbf{V}}) H_2^{1/4}(f_{\mathbf{V}|\mathbf{U}}) H_2^{1/4}(f_{\mathbf{U}}) H_2^{1/4}(f_{\mathbf{V}}),
 \end{aligned}$$

- Equalities iff  $f_{\mathbf{U}, \mathbf{V}}(\mathbf{u}, \mathbf{u}) = f_{\mathbf{U}}(\mathbf{u}) f_{\mathbf{V}}(\mathbf{u})$  and  $f_{\mathbf{U}}(\mathbf{u}) = f_{\mathbf{V}}(\mathbf{u})$
- Right side of (3) minimized when  $f_{\mathbf{U}}$  is uniform over  $S_{n-2}$ .

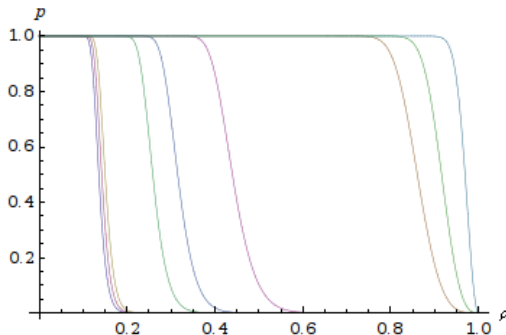
# Implication: phase transition for correlation screening

$M(\rho, n, p)$



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n	550	500	450	150	100	50	10	8	6
$\rho_c$	0.188	0.197	0.207	0.344	0.413	0.559	0.961	0.988	0.9997

Critical threshold approximation:  $\rho_c = \max\{\rho : dE[N]/d\rho = -1\}$

$$\rho_c = \sqrt{1 - c_n(p-1)^{-2/(n-4)}} \quad (3)$$

# Main result: persistent correlation screening

## Proposition

Let the  $n_a \times p$  and  $n_b \times p$  data matrices  $\mathbb{X}^b$  and  $\mathbb{X}^a$  be independent. Let  $\rho_p^a \rightarrow 1$  and  $\rho_p^b \rightarrow 1$  while for  $\gamma = a, b$

$$p^{1/2}(p-1)(1 - (\rho_p^\gamma)^2)^{(n_\gamma-2)/2} \rightarrow d_{n_\gamma}$$

Then

$$\lim_{p \rightarrow \infty} E[N^{a \wedge b}] = \kappa_n^{a \wedge b} \lim_{p \rightarrow \infty} \frac{1}{p} \sum_{i=1}^p J(\overline{f_{\mathbf{U}_i^a, \mathbf{U}_{*-i}^a}}) J(\overline{f_{\mathbf{U}_i^b, \mathbf{U}_{*-i}^b}}), \quad (4)$$

where, for  $\mathbf{U} \in \{\mathbf{U}^a, \mathbf{U}^b\}$ ,

$$\overline{f_{\mathbf{U}_i, \mathbf{U}_{*-i}}(\mathbf{u}, \mathbf{v})} = \frac{1}{p-1} \sum_{j \neq i}^p \left( \frac{1}{2} f_{\mathbf{U}_i, \mathbf{U}_j}(\mathbf{u}, \mathbf{v}) + \frac{1}{2} f_{\mathbf{U}_i, \mathbf{U}_j}(\mathbf{u}, -\mathbf{v}) \right). \quad (5)$$

# Persistent correlation screening: observations

- $\rho_a \rightarrow 1, \rho_b \rightarrow 1$  at slower rates than before.
- When  $J(\overline{f_{\mathbf{U}_i^a, \mathbf{U}_{*i}^a}}), J(\overline{f_{\mathbf{U}_i^b, \mathbf{U}_{*i}^b}})$  are asymptotically *incoherent*

$$\lim_{p \rightarrow \infty} \frac{1}{p} \sum_{i=1}^p J(\overline{f_{\mathbf{U}_i^a, \mathbf{U}_{*i}^a}}) J(\overline{f_{\mathbf{U}_i^b, \mathbf{U}_{*i}^b}}) = J(\overline{f_{\mathbf{U}_\bullet^a, \mathbf{U}_{*\bullet}^a}}) J(\overline{f_{\mathbf{U}_\bullet^b, \mathbf{U}_{*\bullet}^b}})$$

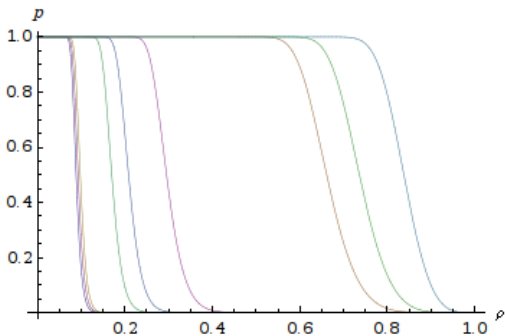
Then, as  $p \rightarrow \infty$ ,

$$E[N^{a \wedge b}] \rightarrow \frac{E[N^a] E[N^b]}{p}$$

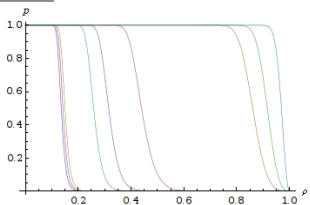
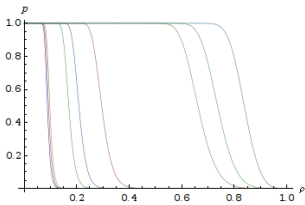
- $p^{-1/2} E[N_a], p^{-1/2} E[N_b]$  converge but  $E[N_a], E[N_b]$  do not.

# Implication: phase transition for persistent correlation screening

$M(\rho, n, p)$



# Phase transitions: correlation vs persistent correlation screening

 $M(\varphi, n, p)$  $M(\varphi, n, p)$ 

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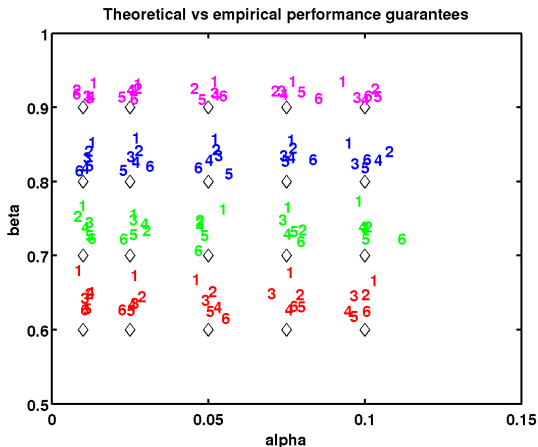


# Application: correlation screening with spike-in

$n \setminus \alpha$	0.010	0.025	0.050	0.075	0.100
10	0.99 \ 0.99	0.99 \ 0.99	0.99 \ 0.99	0.99 \ 0.99	0.99 \ 0.99
15	0.96 \ 0.96	0.96 \ 0.95	0.95 \ 0.95	0.95 \ 0.94	0.95 \ 0.94
20	0.92 \ 0.91	0.91 \ 0.90	0.91 \ 0.89	0.90 \ 0.89	0.90 \ 0.89
25	0.88 \ 0.87	0.87 \ 0.86	0.86 \ 0.85	0.85 \ 0.84	0.85 \ 0.83
30	0.84 \ 0.83	0.83 \ 0.81	0.82 \ 0.80	0.81 \ 0.79	0.81 \ 0.79
35	0.80 \ 0.79	0.79 \ 0.77	0.78 \ 0.76	0.77 \ 0.76	0.77 \ 0.75

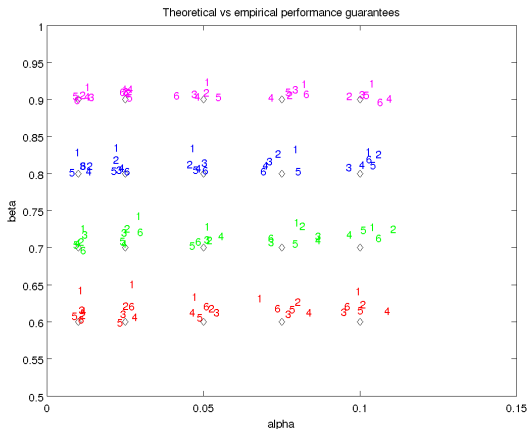
**Table:** Achievable limits in FPR ( $\alpha$ ) for TPR = 0.8 ( $\beta$ ), as function of  $n$ , minimum detectable threshold, and correlation threshold ( $\rho_1 \setminus \rho$ ). To obtain entries  $\rho_1 \setminus \rho$  a Poisson approximation determined  $\rho = \rho(\alpha)$  and a Fisher-Z Gaussian approximation determined  $\rho_1 = \rho_1(\beta)$ . Here  $p = 1000$  on Gaussian sample having diagonal covariance with a spike-in correlated pair.

# Application: correlation screening with spike-in



**Figure:** Comparison between predicted (diamonds) and actual (numbers) operating points  $(\alpha, \beta)$  using Poisson approximation to false positive rate ( $\alpha$ ) and Fisher approximation to false negative rate ( $\beta$ ). Each number is located at an operating point determined by the sample size  $n$  ranging over  $n = 10, 15, 20, 25, 30, 35$ . These numbers are color coded according to the target value of  $\beta$ .

# Application: persistent correlation discovery



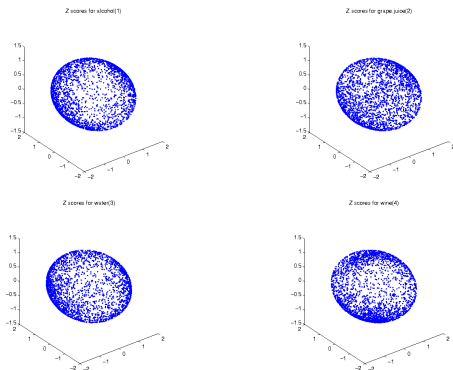
**Figure:** Comparison between predicted (diamonds) and actual (numbers) operating points  $(\alpha, \beta)$  for persistent correlation screening.

# Application: gene expression data

## Beverage Data from Gene Expression Omnibus (GEO) NCBI

- Reference: Florent Baty *etal* (2006) BMC Bioinformatics
- Subjects: 6 individuals at 5 time points (0, 1, 2, 4, 12 hours)
- Treatments: post-baseline intake of
  - $A$ : alcohol ( $n_1 = 20$ )
  - $G$ : grape juice ( $n_2 = 22$ )
  - $H$ : water ( $n_3 = 23$ )
  - $W$ : red wine ( $n_4 = 22$ )
- 87 Affymetrix HU133 Genechip peripheral blood samples
- Each sample contains  $p = 22,283$  gene probes

# Application: observed Z-scores



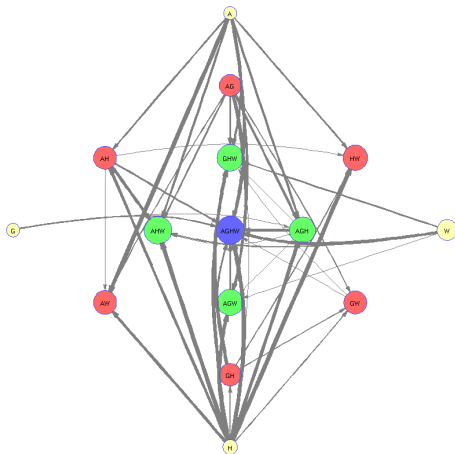
**Figure:** 3 dimensional projections of the Z-scores for the experimental beverage data under each of the treatments A,G,H,W. For visualization the 22,238 variables (gene probes) were downsampled by a factor of 8 and a randomly selected set of four samples in each treatment were used to produce these figures.

# Application: persistent correlation discoveries

$\{A\}, \{G\}, \{H\}, \{W\}$		42	50	82	424	
$\{A, G\}, \{A, H\}, \{A, W\}, \{G, H\}, \{G, W\}, \{H, W\}$	493	748	1069	677	864	1445
$\{G, H, W\}, \{A, H, W\}, \{A, G, W\}, \{A, G, H\}$		2242	2530	1893	1690	
$\{A, G, H, W\}$			3313			

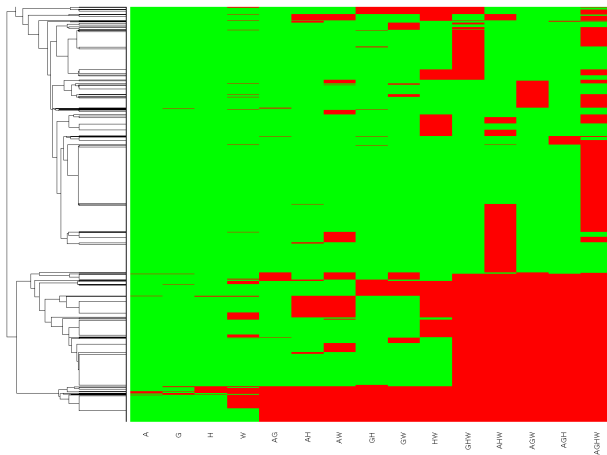
**Table:** Number of genes discovered by auto-screening (top row) and persistency screening (lower three rows) for various combinations of treatments in the experimental data. Auto-screening threshold determined using Poisson approximation to Type I error of level  $10^{-5}$ .

# Application: set-inclusion diagram



**Figure:** Set-inclusion graph between genes discovered by correlation screening in various combinations of treatments. Size of node is proportional to the log of number of associated correlation screening discoveries given in Table 2. A directed edge from node  $i$  to node  $j$  exists if at least 90% of the genes discovered in node  $i$  are also discovered in node  $j$  and the thickest edges indicate 100% set inclusion. The asymmetry of diagram indicates that treatments have different effects on gene expression.

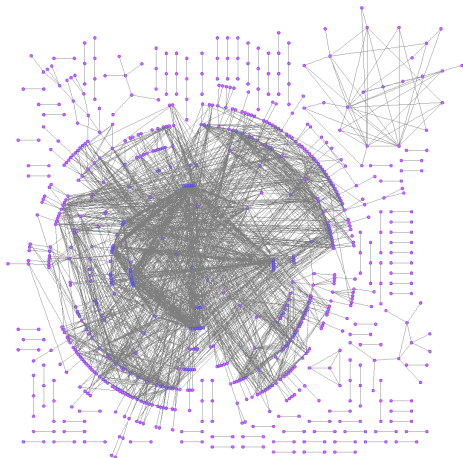
# Application: persistent covariance network



**Figure:** Heatmap of 4444 genes discovered in at least one of the set inclusion tests shown in Table 2.



# Application: persistent covariance network



**Figure:** 774 gene subnetwork of the 3313 gene persistent-correlation network across all four treatments corresponding to the last row of Table 2. Two nodes in this network are linked by an edge if for all 4 treatments their sample correlation is above the  $10^{-5}$  FWER correlation-screening threshold.

# Outline

- 1 Motivation
- 2 Theory
- 3 Application
- 4 Conclusions**

# Conclusions

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  - Screening affected by phase transition as threshold decreases
  - Large  $p$  expressions for critical PT threshold  $\rho_c$  are available
  - Effect of pairwise dependence manifested through Hellinger divergence

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  - Geometry of unit sphere
- Persistence: Strongest specialists are not strongest generalists